

Currency Factors

Arash Aloosh

Neoma Business School

arash.aloosh@neoma-bs.fr

Geert Bekaert

Columbia Business School

gb241@columbia.edu

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Abstract:

We examine the ability of existing and new factor models to explain the comovements of G10-currency changes. Extant currency factors include the carry, volatility, value, and momentum factors. Using a new clustering technique, we find a clear two-block structure in currency comovements with the first block containing mostly the dollar currencies, and the other the European currencies. A factor model incorporating this “clustering” factor and two additional factors, a commodity currency factor and a “world” factor based on trading volumes, fits all bilateral exchange rates well, whatever the currency perspective. In particular, it explains on average about 60% of currency variation and generates a root mean squared error relative to sample correlations of only 0.11. The model also explains a considerable fraction of the variation in emerging market currencies.

Keywords:

Currency Factors; Currency Comovements; Clustering; Factor Models.

Introduction

According to recent Bank of International Settlement Surveys, more than half the trading volume in foreign exchange arises from trades with “financial” customers, institutional investors, mutual funds, hedge funds and other portfolio managers (Rime and Schrimpf, 2013), a phenomenon that mostly reflects increasing globalization over time. While the need for foreign exchange (FX) transactions often arises merely as a by-product of buying or selling international securities, increasingly, there are also profit-seekers in currency markets, seeking to profit from selling and buying currencies.¹ With active currency management becoming more commonplace, the need for models to explain the risks and comovements of currencies has increased. There is also renewed interest in the pricing of currency risk in international equities (see e.g., Brusa, Ramadorai and Verdelhan, 2015), but standard theory (see the famous Adler and Dumas, 1983 survey) suggests that all currency risks are priced for each equity market. A parsimonious currency factor model may therefore help the implementation of international equity pricing models.

The academic literature so far has focused almost exclusively on detecting currency factors that generate attractive return profiles. Two important currency factors include the carry factor of Lustig, Roussanov and Verdelhan (2011) and the global FX volatility factor of Menkhoff, Sarno, Schmeling, and Schrimpf (2012a). Other extant currency factors include currency-value (see e.g., Menkhoff, Sarno, Schmeling, and Schrimpf, 2016) and currency-momentum factors (see e.g., Menkhoff, Sarno, Schmeling, and Schrimpf, 2012b, and Burnside, Eichenbaum, and Rebelo, 2011), which also feature in practitioner indices created by Deutsche Bank. Practitioners also recognize that there is a commodity factor in currencies, and the Australian and Canadian dollar are typically categorized as “commodity currencies,” see Chen and Rogoff (2003) and Ready, Roussanov, and Ward (2013). Just as the Fama–French model (1996) for equities is also a good risk model to explain equity return comovements (see Bekaert, Hodrick, Zhang, 2009 and Hou, Karolyi and Kho, 2011), it may be that these factors are effective in explaining currency comovements.

¹ Hafeez (2007), then the Global Head of FX Strategy at Deutsche Bank, estimated that between 5% and 25% of the trading volume in the FX market can be categorized as profit-seeking.

Developing an adequate factor model for currency movements raises special issues however. If we take the dollar as the numeraire currency, a factor model that explains the bilateral dollar movements perfectly, will, by triangular arbitrage, also fit other bilateral exchange rates perfectly, whatever the perspective. However, if the fit is imperfect, a good dollar model may be a poor yen model and vice versa. This tension is obvious in the recent work of Verdelhan (2016). Verdelhan (2016) introduces the dollar basket factor, the average appreciation of the U.S. dollar relative to a basket of currencies, as a common factor and shows that it has a very strong explanatory power (high adjusted R^2 s) for the contemporaneous bilateral exchange rate changes w.r.t. the U.S. dollar. However, its explanatory power is quite poor for other currency perspectives. To address this issue, Verdelhan (2016) introduces a different common factor for the exchange rates denominated in other currencies, for example a pound factor for the British pound rates.

In this paper, we set out to examine various factor models to explain currency comovements and document their fit with the data from a *global* perspective. That is, we attempt to identify a factor model that works well whatever the currency perspective is. To facilitate a global perspective on currency comovements, we introduce the concept of a “currency basket.” The currency basket simply averages all bilateral currency changes relative to one particular currency. As we show formally, by analyzing 10 currency baskets for the G10 currencies, we span all possible bilateral currency movements. We then contrast the explanatory power of the extant risk factors mentioned previously with the explanatory power of various new factors.

Most importantly, we use a new clustering technique to introduce several new currency factors. When selecting two clusters, a very clear factor structure emerges, with the dollar currencies (Australian, Canadian, New Zealand and US) and the Japanese yen in one block and the European currencies in the other. When using three clusters, a commodity type currency factor also emerges. Combining these statistical factors with a “market” factor, based on currency trading volumes, and a commodity currency factor, we propose several parsimonious factor models and run a horse race versus models incorporating the existing factors.

Among the extant currency factors, the carry and value factors exhibit the highest explanatory power for currency variation. This is not surprising because both factors are relatively highly correlated with the first principal component in bilateral currency rates. However, a new

parsimonious factor model incorporating the two-block clustering factor, a commodity factor and the market factor easily beats factor models created from extant risk factors, even models that feature double as many factors. The new factor model explains on average about 60% of the variation in changes in currency basket values. Moreover, the Root Mean Squared Error (RMSE) relative to sample correlations is only about 0.11, which is statistically significantly better than any model based on extant risk factors.

The remainder of the article is organized as follows. In Section I, we describe our methodology and introduce the concept of a currency basket. Section II explains our clustering technique and introduces a new factor model for currency returns. Section III examines the explanatory power of the standard currency factor models for currency comovements. Section IV runs a horse race of a variety of different factor models, using primarily the RMSE for correlations as the metric. In Section V, we investigate recent factor models, including the factor model of Verdelhan (2016) and re-interpret the results in Lustig and Richmond (2016), who uncover a gravity equation in the factor structure of exchange rates. We also relate our work to the tradeable currency factors marketed by Deutsche Bank. In Section VI, we examine the explanatory power of our new factor model for emerging market currencies showing it to explain a smaller portion of their variation but fit comovements about as well as for developed currencies. Section VII concludes.

I. Explaining Currency Comovements

We study the G10 currencies—AUD, CAD, CHF, EUR, JPY, NOK, SEK, NZD, GBP, and USD. We use end-of-month mid spot rates from Barclays Bank International (BBI) and WM/Reuters (WMR) that are available via Thomson Reuters Datastream. The time period is from January 1973 to December 2015.

I.1. Bilateral Correlations

To set the stage, we first focus on bilateral correlations from two perspectives. In Table 1, Panels I and II show correlation matrices. Panel I takes the perspective of a U.S. dollar investor, and reproduces the full sample correlation matrix for all currency pairs relative to the dollar expressed

in foreign currencies per dollar; e.g., CAD/USD is the amount of Canadian dollar equivalent to one U.S. dollar); Panel II takes the perspective of a yen investor. Panel III provides the differences between correlations in Panels I and II. The correlations in Panel I show that the “*dollar currencies*” (USD, AUD, CAD, and NZD) are more correlated with each other and less correlated with the other currencies, e.g., in the second column log changes in the AUD/USD are 55% and 69% correlated with log changes in the CAD/USD and NZD/USD respectively, while correlations between log changes in the AUD/USD and the remaining currencies w.r.t. the USD vary from only 17% to 44%. In addition, the other currencies (EUR, NOK, SEK, CHF, and GBP) are more correlated with each other and less correlated with the other currencies, e.g., in the ninth row the correlation between log changes in the GBP/USD rate ($\Delta_{\text{GBP,USD}}$) and $\Delta_{\text{EUR,USD}}$, $\Delta_{\text{NOK,USD}}$, $\Delta_{\text{SEK,USD}}$, and $\Delta_{\text{CHF,USD}}$ is 69%, 66%, 63%, and 60% respectively, while the correlations between log changes in the GBP/USD and the dollar currencies w.r.t. the USD vary from only 31% to 42%. Log changes in the JPY/USD rate ($\Delta_{\text{JPY,USD}}$) have the lowest correlations with other currencies on average.

The correlations in Panel II are similar to those in Panel I, the dollar currencies are more correlated with each other and less correlated with the other currencies (EUR, NOK, SEK, CHF, and GBP) while the other currencies are more correlated with each other and less correlated with the dollar currencies. The correlation differences in Panel III are mostly negative and often quite substantially so. In other words, most of pairwise currency correlations are higher in Panel II compared to those in Panel I. In particular, if we contrast the correlation between the AUD, CAD, and NZD with dollar yen rates either from the yen (Panel II) and the dollar (Panel I) perspective, the correlations are respectively 46%, 82%, and 32% higher from the yen perspective.

These numbers demonstrate two facts. First, the correlation structure among currencies is very numeraire dependent. A factor model that fits “dollar based” correlations well, may not fit “yen based” correlations well. Second, there appears to be, nonetheless, an overall factor structure in currencies with certain currencies correlating more with one another irrespective of the base currency. We seek to find factor models for currencies that maximize overall fit, across all base currencies.

[Table 1]

I.2. Currency Baskets

Explaining currency comovements globally is non-trivial. The 10 currencies imply 45 different currency pairs, which are linearly dependent through the triangular arbitrage relation. To resolve this problem, we introduce the concept of a *currency basket*.

A currency-basket-factor is an equally-weighted average appreciation of one currency relative to a basket of all currencies in our sample. In other words, the currency i -basket-factor is calculated as

$$CB_i = \frac{1}{9} \sum_j \Delta s_{j,i} \quad (1)$$

where $\Delta s_{j,i}$ is the log spot rate change of currency i w.r.t. currency j that is, the (logarithm) change in the value of currency i relative to currency j . For example, the U.S. dollar-basket-factor denoted by CB_{USD} is an equally-weighted average of log changes in the value of the U.S. dollar w.r.t. AUD, CAD, CHF, EUR, JPY, NOK, SEK, NZD, and GBP.

Under the absence of triangular arbitrage, we can replicate all bilateral rates by having only $N-1$ non-repeated exchange rates. It should therefore not be surprising that our 10 currency baskets span all bilateral rates. Because the concept of the currency basket is essential to this article, we show this spanning property in some detail.

Triangular arbitrage implies that:

$$\Delta s_{k,i} = \Delta s_{j,i} + \Delta s_{k,j} \quad \forall j. \quad (2)$$

In the absence of arbitrage, this equation holds for any third currency. Therefore, we can add up "n" of those triangular equations for a basket of third currencies to find the relation between the log appreciation of a bilateral exchange rate and the currency basket factors. From Equation (2), we have;

$$\Delta s_{k,i} = \frac{1}{n} n (\Delta s_{k,i}) = \frac{1}{n} \sum_{j=1}^n [\Delta s_{j,i} + \Delta s_{k,j}] = \frac{1}{n} [\sum_{j=1}^n \Delta s_{j,i}] - \frac{1}{n} [\sum_{j=1}^n \Delta s_{j,k}] \quad (3)$$

If there are "n" currencies, there will be "n-1" exchange rates. Therefore,

$$\Delta s_{k,i} = \frac{1}{n} [\sum_{j \neq i}^{n-1} \Delta s_{j,i} + \Delta s_{i,i}] - \frac{1}{n} [\sum_{j \neq k}^{n-1} \Delta s_{j,k} + \Delta s_{k,k}] \quad (4)$$

As $\Delta s_{i,i} = \Delta s_{k,k} = 0$, we get

$$\Delta s_{k,i} = \frac{n-1}{n} \frac{1}{n-1} [\sum_{j \neq i}^{n-1} \Delta s_{j,i}] - \frac{n-1}{n} \frac{1}{n-1} [\sum_{j \neq k}^{n-1} \Delta s_{j,k}] \quad (5)$$

$$\Delta s_{k,i} = \frac{n-1}{n} CB_i - \frac{n-1}{n} CB_k \quad (6)$$

Equation (6) simply shows that the appreciation of currency k w.r.t. currency i ($\Delta s_{k,i}$) is spanned by the average appreciation of a basket of currencies w.r.t. currency i (CB_i) minus the average appreciation of a basket of currencies w.r.t. currency k (CB_k). Empirically, we have nine bilateral exchange rates among G10 currencies to estimate CBs. Therefore, using the definition of currency baskets in Equation (1), it follows:

$$\Delta s_{k,i} = \frac{9}{10} CB_i - \frac{9}{10} CB_k \quad (7)$$

Table 2 reports summary statistics on currency basket factors. Over the sample period, CB_{SEK} has the highest annualized depreciation rate of 1.2 %, whereas CB_{CHF} has the highest annualized appreciation rate of 3.7 %. Annualized volatilities range between 5.8% for CB_{EUR} and 10.5 % for CB_{JPY} . Importantly, the returns of the currency baskets show little serial correlation with the first-order autocorrelations never higher than 0.10 in absolute value.

[Table 2]

Note that the U.S. dollar basket factor (CB_{USD}) corresponds to the dollar factor introduced in Lustig, Roussanov and Verdelhan (2011). It is likely that this factor explains exchange rate changes from the dollar perspective well, but how well does it explain currency changes from other perspectives? To examine this, we construct currency baskets for each currency and test their explanatory power (average adjusted R-squares) for bilateral exchange rate changes. Table 3 compares the contemporaneous explanatory power of currency-basket factors for bilateral exchange rates from different currency perspectives. That is, we run regressions of the form,

$$\Delta s_{j,i} = a_j + b_j CB_k + e_{j,i}, \text{ for all currency perspectives } i \text{ and currency baskets } k. \quad (8)$$

Not surprisingly, each currency-basket factor has the highest explanatory power for its own bilateral rates. For example, the second column of Table 3 shows that the U.S. dollar-basket factor (CB_{USD}) explains 55.4% of the variation in the bilateral exchange rates against the U.S. dollar (USD rates), which is analogous to the explanatory power documented in Verdelhan (2016). However, the explanatory power of CB_{USD} is low for the other bilateral rates, varying from 7.7% for JPY rates to 23.4% for CAD rates. More interestingly, although the CB_{USD} factor explains on average 16.7% of all exchange rate variation, the Swiss frank basket (CB_{CHF}) explains on average 19.8 % of all exchange rate variation, constituting the highest explanatory power among the G10 currency-basket factors. The last row represents the average off-diagonal adjusted R-square. CB_{AUD} , CB_{EUR} , and CB_{CHF} deliver the highest explanatory power among the G10 currency-basket factors. Clearly, by triangular arbitrage, there is dependence among these rates. Yet, Table 3 shows that it is not obvious which combination of currencies would capture correlations well for all currency perspectives. By focusing on currency baskets, we collapse a total of 45 different bilateral rates that are codependent into 10 manageable baskets.

In the right two columns, we use the top three or five currency baskets in each row to explain bilateral currency movements and report the adjusted R^2 . These always include the own basket. By the spanning argument that we discussed earlier, the R^2 rapidly increases and reaches on average 82% with 5 baskets. For the remainder of our paper, we examine which factor models best described the correlation structure of the currency baskets. These models will then automatically also describe comovements between any bilateral rates.

[Table 3]

I.3. Numeraire issues

All our computations have made use of logarithmic exchange rate changes. This generates two unit issues. First, investors care about returns not just currency changes. Second, by considering various nominal currency baskets simultaneously, we aggregate economic variables expressed in different currencies. Both issues are in fact immaterial given our objective of creating a factor model that works from all currency perspectives. The main reason for this is that the variability of currency changes is almost an order of magnitude larger than the variation of interest and

inflation differentials and thus nominal currency changes are the main driver of currency comovements.

To verify this, we compute excess bilateral exchange rate returns as well as two real concepts: bilateral real exchange rate changes and real foreign exchange returns. The excess returns are calculated as the one-month exchange rate changes plus the monthly interest rate differentials. The real exchange rate changes are calculated as the one-month exchange rate changes plus monthly inflation rate differentials. Real returns are computed as the nominal exchange rate changes plus the foreign interest rate deflated by domestic inflation. The interest rate and inflation data are non-seasonally adjusted and available on Datastream.

We compute the equivalent currency baskets in excess return space and correlate them with the currency baskets using currency changes. The lowest correlation is observed for the GBP currency basket, and the correlation is 99.7%.

Excess returns are still priced by different nominal pricing kernels in different currencies. In an integrated economy, a world pricing kernel should price real returns in various countries; alternatively, if purchasing power parity holds, the real return from investing in any country would be equalized whatever the numeraire perspective. We therefore also formulate the currency baskets in real return space. Here, the correlations between “real” currency baskets (real currency changes) and our nominal exchange rate ones vary between 95.6% for the EUR currency and 99.1% for the AUD currency. For actual real returns, the correlations vary between 98.68% for the USD perspective and 99.92% for the CAD perspective.

I.4. Factor models

We examine a variety of linear factor models to maximize “global” fit:

$$CB_j = a_j + \beta'_j \times F + e_j, \quad (9)$$

where F is a set of factors and β_j the vector of factor exposures. To examine and compare fit, we focus on a number of different statistics.

First, we simply examine the significance of the betas in Equation (9) and calculate a global R-square as the equally weighted average of the R^2 's for each CB_j. We redid all computations with R^2 's based on trading volumes (see below) as well. Because our conclusions are robust to using such alternative global R^2 's, we do not discuss them further (results are available upon request).

Second, we examine how well the various factor models explain the comovement structure present in exchange rates. Let $\widehat{\sigma}_{i,j}$ be the empirical covariance between CB_i and CB_j. Using the implied covariances from a particular factor model, we can compare the correlation fit of different factor models. The covariance matrix produced by a particular factor model with factor covariance matrix V_F is, as usual,

$$\text{Cov}_F = \beta_F' V_F \beta_F \quad (10)$$

where β_F is the 10xK matrix of factor loadings, K the number of factors, and Cov_F the model-implied covariance matrix for the currency basket factors.

We then compute the correlation fit for model F as

$$\text{RMSE}_F = \sqrt{\frac{1}{45} \sum_j \sum_{i \neq j} (\hat{\rho}_{i,j} - \rho_{i,j})^2} \quad (11)$$

where $\rho_{i,j}$ is the model implied correlation between currency i and j; $\rho_{i,j} = \frac{\beta_{F,i} V_F \beta_{F,j}'}{\hat{\sigma}_i \hat{\sigma}_j}$, $\hat{\sigma}_i$ and $\hat{\sigma}_j$ are the sample variances, $\hat{\rho}_{i,j}$ is the sample correlation, and RMSE stands for root mean squared error.

Finally, we compare the correlation fit of various currency factor models, seeking models that minimize the RMSE for correlation. To account for sampling error in those computations, we conduct a bootstrap exercise, in which we bootstrap the 10 currency baskets with replacement. The Bootstrap creates artificial samples of equal length to our sample by randomly selecting and concatenating blocks of 6 months of currency basket changes. The contemporaneous correlation structure is therefore preserved. For each random sample, we estimate the correlation matrix as well as the factor model. Then, we use the factors exposures to compute model-implied correlations and finally the RMSEs. We use 1000 replications.

II. A New Factor Model for Currency Returns

Here we propose a new currency factor model that incorporates a statistical factor, a factor based on trading volumes (akin to the market model often used in equity trading) and the one currency factor that is perhaps most often referred to in practitioner's circles, a commodity currency factor. Importantly, an intuitive clustering technique uncovers a very prevalent two-block factor structure in currencies, which is the main focus of this section.

II.1. Cluster Analysis

The correlations in Table 1 are suggestive that there may be a two or three factor structure in currencies. Dollar rates seem highly correlated, as do rates within continental and Scandinavian Europe. To investigate this formally, we use a clustering technique introduced by Ormerod and Mounfield (2000), and used to investigate the clustering of currencies just before the Euro was introduced. Ahn, Conrad, and Dittmar (2009) apply the algorithm in a stock portfolio formation context to create "basis assets." They show that their algorithm produces cluster portfolios that are correlated with the standard firm characteristics, display significant dispersion in returns, and generate a relatively well-conditioned return covariance matrix.

The algorithm starts by defining a distance measure, which is a negative function of correlation:

$$d_{ij} = \sqrt{2 * (1 - \rho_{ij})}, \quad (12)$$

where ρ_{ij} denotes the sample correlation between currency baskets i and j , CB_i and CB_j , respectively. Perfectly positively correlated currency baskets have the minimum distance of 0 whereas perfectly negatively correlated currency baskets have the maximum distance of 2. Note that $d_{ii} = 0$. The clustering algorithm then creates clusters aiming to maximize within-group correlation and minimize across-group correlations.

An obvious way to use the distance concept to cluster currencies into N clusters, is to find the combination of currencies that minimizes the total distance between currency baskets within a cluster. This absolute clustering algorithm therefore minimizes,

$$SD(N) = \sum_{k=1}^N \sum_{i,j} d_{ij}, \quad (i,j \in k^{\text{th}} \text{ cluster}), \quad (13)$$

where k indexes a cluster of currencies, N is the number of clusters and SD stands for sum of distances between all members of the cluster.

In other words, to cluster currencies, we first consider all possible allocations of G10 currency baskets in N different clusters ($1 \leq N \leq 10$) and calculate their in-cluster distance as the sum of distances among all members of each cluster. Then, we calculate the total distance as the sum of all in-cluster distances for each possible allocation. The currency allocation that minimizes the total distance for each N constitutes the optimal clustering of the G10 currency baskets in N clusters.

Given our limited set of currencies, we can easily consider all possible combinations of currency allocations for a given number of N clusters. However, in the aforementioned papers, the authors applied a sequential clustering procedure. At the beginning of such process, each currency basket is placed into its own individual cluster; thus, we start with N clusters for N currency baskets; therefore the sum of distances is zero (by definition). The algorithm then proceeds sequentially by optimally joining the individual currency baskets, and later, clusters of currency baskets. That is, for every possible combination of smaller clusters i and j (CL_i and CL_j), the algorithm computes:

$$IL_{ij} = SD(CL_{ij}) - [SD(CL_i) + SD(CL_j)], \quad (14)$$

where $SD(CL_{ij})$ is the sum of distances among currency baskets in the new aggregate cluster, and $SD(CL_i)$ and $SD(CL_j)$ are the sum of distances among currency baskets in the smaller clusters i and j , respectively. The “information loss” IL_{ij} is then summed over all cluster aggregations.

The cluster aggregation that minimizes the increase in the total distance is selected. We can repeat this process until currencies are clustered into two blocks.

II.2. Optimal Currency Clusters

To gain some intuition regarding the methodology, Table 4 reports the results of clustering G10 currency baskets in N clusters, with N varying from 0 to 10, contrasting the optimal clusters

generated by both the absolute algorithm and sequential algorithms (in Panels I and II, respectively). The end points are trivial: for 10 clusters, each currency basket is in its own cluster and the total distance is zero; for the case of 1 cluster, the distance reflects the average correlation of all currency baskets. For 9 to 5 clusters, the sequential and absolute cluster algorithms produce the same clusters, because clusters of more than 2 currencies generate low within correlations relative to pairing up currencies, which appears the optimal recombination procedure suggested by the sequential procedure. The optimal pairing of currencies is largely achieved along regional lines, starting with CB_{CAD} and CB_{USD} , then CB_{CHF} and CB_{EUR} , CB_{NZD} and CB_{AUD} , CB_{NOK} and CB_{SEK} , and finally CB_{JPY} and CB_{GBP} .

[Table 4]

For 2, 3 or 4 currency-clusters, the sequential clustering algorithm starts to underperform the absolute algorithm, which can consider all possible allocations of G10 currencies in K clusters ($K=2, 3$, or 4) while the sequential algorithm can only merge two clusters among $K+1$ clusters given in a previous step. For example, to cluster the G10 currency baskets in 3 clusters, the absolute algorithm produces (CB_{AUD} , CB_{CAD} , and CB_{NZD}), (CB_{CHF} , CB_{EUR} , CB_{NOK} , and CB_{SEK}) and (CB_{USD} , CB_{JPY} , and CB_{GBP}) with a total distance of 14.830 (average within-cluster correlation of 0.577), while the sequential algorithm produces (CB_{USD} , CB_{AUD} , CB_{CAD} , and CB_{NZD}), (CB_{CHF} , CB_{EUR} , CB_{NOK} , and CB_{SEK}) and (CB_{JPY} and CB_{GBP}) with a total distance of 15.957 (average within-cluster correlation of 0.565). To reach an optimal three clusters, the sequential clustering algorithm only considers six possible mergers among the 4 optimal clusters in the previous step, namely (CB_{USD} and CB_{AUD}), (CB_{CAD} and CB_{NZD}), (CB_{CHF} , CB_{EUR} , CB_{NOK} , and CB_{SEK}) and (CB_{JPY} , and CB_{GBP}).

In this article, we focus on the clusters generated by the absolute algorithm. Results with clusters from the sequential algorithm are available upon request, but not surprisingly, the factor model resulting from the absolute algorithm performs better. Figure 1 shows how the clustering algorithm lowers the total distance for all clusters, $N=1$ through 10. Because we seek to construct a parsimonious factor model, let's focus on $N=2$. For two clusters, the worst grouping generates a total distance of 53.4, which corresponds to an average within-cluster correlation of 0.125. When we use the algorithm to minimize the distance, it more than halves to 26.2, and the average within-cluster correlation is much higher at 0.416.

Optimal clustering for two clusters suggests that the dollar currencies plus the Japanese yen (CB_{USD} , CB_{AUD} , CB_{CAD} , CB_{NZD} , and CB_{JPY}) are clustered in the same block, and the European currencies (CB_{CHF} , CB_{EUR} , CB_{NOK} , CB_{SEK} , and CB_{GBP}) are clustered in the other block. The currency basket correlations reported in Table 5 confirm that currency baskets are more positively correlated within these blocks and more negatively correlated across the blocks.

[Table 5]

Note that the first block involves all “dollar” currencies plus the Japanese yen, whereas the other block involves all European currencies. Therefore, the countries in each currency block share commonality in language, border, legal origin, culture, and resources, or have colonial linkages, features stressed in recent work by Lustig and Richmond (2015). We link our work to theirs explicitly in Section V.

Based on these currency blocks, we introduce a currency factor (CF_{abs}), as the sum of the dollar currency basket returns plus the Japanese yen basket return, as follows:

$$CF_{abs} = CB_{USD} + CB_{AUD} + CB_{CAD} + CB_{NZD} + CB_{JPY} \quad (15)$$

Because each currency pair appears in two currency baskets with opposite signs, the sum of all currency baskets equals zero; that is,

$$CB_{USD} + CB_{AUD} + CB_{CAD} + CB_{CHF} + CB_{EUR} + CB_{JPY} + CB_{NOK} + CB_{SEK} + CB_{NZD} + CB_{GBP} = 0. \quad (16)$$

Therefore, the sums of currency basket changes in the two blocks are perfectly negatively correlated. In addition, using the definition of a currency basket in Equation (1), we can show that:

$$CB_{USD} + CB_{AUD} + CB_{CAD} + CB_{NZD} + CB_{JPY} = (-\Delta s_{AUD,USD} - \Delta s_{CAD,USD} - \Delta s_{JPY,USD} - \Delta s_{NZD,USD} + \Delta s_{EUR,USD} + \Delta s_{NOK,USD} + \Delta s_{SEK,USD} + \Delta s_{CHF,USD} + \Delta s_{GBP,USD}) \times 5/9 \quad (17)$$

Equation (17) shows how the CF_{abs} factor represents an investment strategy of longing dollar currencies as well as Japanese yen and shorting European currencies.

II.3. A New Factor Model

The bottom row of Table 5 presents the correlations between our CF_{abs} factor and the currency-baskets. Not surprisingly, the CF_{abs} is positively correlated with CB_{USD} , CB_{AUD} , CB_{CAD} , CB_{NZD} , and CB_{JPY} , and negatively correlated with CB_{CHF} , CB_{EUR} , CB_{NOK} , CB_{SEK} , and CB_{GBP} . The CF_{abs} is most highly correlated with CB_{AUD} , CB_{CAD} , CB_{CHF} , and CB_{EUR} . Its absolute correlation with currency baskets varies from 19% to 75%, averaging 54.2%, making it an excellent candidate as a currency factor.

In Table 6, the first column shows regression coefficients from regressing the currency baskets onto the CF_{abs} factor. The coefficients are highly statistically significant for all currency baskets, with R^2 's ranging between 14% and 43%. The Online Appendix shows that the CF_{abs} factor is also highly correlated with the bilateral rates directly, with its explanatory power better than any “off-diagonal” currency basket.

We consider two avenues to come up with a parsimonious model for currency comovements. First, we create three clusters instead of two clusters. Going back to Table 4, this yields two currency factors (as the third one is co-linear with the other two),

$$CF_{31} = CB_{USD} + CB_{GBP} + CB_{JPY} \quad (18)$$

$$CF_{32} = CB_{AUD} + CB_{CAD} + CB_{NZD} \quad (19)$$

The first factor combines the USD with the British pound and the Yen; whereas the second factor combines all the other dollar rates. Interestingly, the latter factor would be close to what practitioners would dub a commodity currency basket (which would also involve the NOK). In Table 6, we regress the currency baskets on both factors, showing that both are highly statistically significant for all currency baskets. The R^2 's now range between 35% and 59%, averaging 48%.

[Table 6]

Second, we continue to use the CF_{abs} factor but add two “economic” factors. The first is the commodity factor (CF_{com}), computed as the sum of commodity-driven currency basket returns including CB_{AUD} , CB_{CAD} , CB_{NZD} , and CB_{NOK} :

$$CF_{com} = CB_{AUD} + CB_{CAD} + CB_{NZD} + CB_{NOK} \quad (20)$$

The second is the market factor (CF_{TW}) computed as the trading-volume weighted average of all G10 currency basket returns:

$$CF_{TW} = \sum_i^{10} w_i CB_i, \quad i \in \{\text{G10 currencies}\}. \quad (21)$$

where i indexes the G10 currencies and w_i represent the trading-volume weights reported by the Bank for International Settlements (BIS) every three years from 1998 to 2013. We fix the weights before 1998 at the 1998 weights. In addition, the BIS weights include non-G10 currencies and add up to 200%, because each currency trade is counted twice for both trading parties. Thus, we calculate a new weight for each G10 currency as its BIS weight divided by the sum of all G10 currencies' BIS weights (see Appendix A, Table A1). The weights are highest for the dollar (around 50%), followed by the euro (around 20%) and the yen (around 10%). The explanatory power of these two factors in isolation is quite substantial but somewhat lower than that of the CF_{abs} factor (full results are relegated to the Online Appendix).

In Table 6, Panel III, we report the results from a regression of the currency baskets onto this first candidate factor model with three factors, including CF_{abs} , CF_{com} and CF_{TW} . The bulk of the individual coefficients is highly statistically significant with only 3 out of 30 not significant at the 10% level. The R^2 's now range from 41% to 81% and average 58%. As a second candidate model, we supplement the CF_{31} and CF_{32} factors with the CF_{TW} factor. Recall that the CF_{32} factor is almost a commodity factor so adding the trade-weighted market factor makes the most sense. The last block in Table 6 shows the explanatory power of this candidate factor model. The model's explanatory power is equally impressive with the coefficients mostly highly statistically significant and only 4 coefficients not significant at the 10% level. The R^2 's range from 46% to 70%, but also average 58%.

III. Standard Currency Factors

The extant currency literature has spawned a number of factors inspired by risk considerations (e.g., carry), economic value (PPP based factors) or trading models (momentum). Here, we provide a new perspective by examining the ability of these currency factors to explain the correlation structure among currency changes.

III.1. The Factors

Currency Carry Factor

We use the currency carry factor of Lustig, Roussanov, Verdelhan (2011) representing the return of going long a portfolio of high interest rate (developed) currencies and going short a portfolio of low interest rate (developed) currencies.² This monthly factor is available on Verdelhan's website.

Currency Volatility

Menkhoff, Sarno, Schmeling, and Schrimpf (2012a) suggest that global foreign exchange volatility is important in cross-sectional tests of currency portfolios sorted by the forward discount. Similar to Menkhoff, *et. al* (2012a), we calculate the absolute daily log exchange rate change ($|\Delta s|$) for each currency k on each day τ in our sample (we take all currency perspectives and not only the U.S. dollar). We then average over all currencies available on any given day and average the daily values within each month. Our global FX volatility proxy in month t is therefore given by

$$\sigma_t^{FX} = \frac{1}{T_t} \sum_{\tau \in T_t} \left[\sum_{k \in K_\tau} \left(\frac{|\Delta s|}{K_\tau} \right) \right], \quad (22)$$

where K_τ denotes the number of available currencies on day τ and T_t denotes the total number of trading days in month t . The variable σ_t^{FX} represents our currency volatility factor. Note that the currency-volatility factor does not represent a return or currency change, thus its units are not as easy to interpret as those of the other factors.

Currency Value

By far the most popular fundamental exchange rate model is Purchasing Power Parity (PPP) (see e.g., Abuaf and Jorion (1990) and the survey of Lothian and Taylor (1996)). If exchange rates revert back to their long-term PPP values (see e.g., Mark (1995) for empirical evidence), similar deviations from PPP can be a source of currency comovements.

² Their dataset contains the currencies of 15 developed countries: Australia, Belgium, Canada, Denmark, euro area, France, Germany, Italy, Japan, Netherlands, New Zealand, Norway, Sweden, Switzerland, and the United Kingdom.

We create a PPP factor return in three steps. First, we obtain PPPs for 28 countries from the OECD for the period from January 1973 to December 2015.³ These PPPs reflect annual averages of monthly values and vary over the year. The OECD constructs PPPs for detailed items that are part of GDP (e.g., Fish) and aggregates them using relative expenditures.

Second, for each month and each currency, we create a currency value index as a currency's nominal exchange rate divided by its PPP last year. For example, the value ratio for GBP/USD is

$\frac{S_t^{\text{GBP/USD}}}{PPP_{t-12}^{\text{GBP/USD}}}$, where $S_t^{\text{GBP/USD}}$ is the average daily GBP/USD spot rate over the last three months and $PPP_{t-12}^{\text{GBP/USD}}$ is the average annual PPP for GBP/USD over the last year.

We then sort the currencies into three groups based on those valuation ratios. Finally, our value factor represents the returns (including interest rate differentials) on the portfolio of under-valued currencies minus the returns on the portfolio of the over-valued currencies. The portfolios are rebalanced every three months.

Currency Momentum

We create a momentum factor return in two steps. First, we obtain spot exchange rates for 28 countries (see Footnote 4) for the period from January 1973 to December 2015. Second, we sort the available currencies into three portfolios based on their exchange rate changes over the last month. Our momentum factor represents the equally-weighted average return on the portfolio of the highest lagged return currencies minus the return on the portfolio of the lowest lagged return currencies. At the end of each month, we rebalance the portfolio. This factor is similar to the 1-month momentum factor of Menkhoff, Sarno, Schmeling, and Schrimpf (2012b).

Commodity Factor

The values of the commodity currencies (AUD, CAD, NOK, and NZD) are correlated with commodity prices. In addition, changes in commodity prices have predictive power for currency carry returns (Bakshi and Panayotov (2013)) as well as for bilateral forex returns (Aloosh (2012)).

³ Our dataset to construct the value factor includes currencies from 28 countries: Australia, Austria, Belgium, Canada, Czech Republic, Denmark, Europe, Finland, France, Germany, Hungary, Iceland, Ireland, Italy, Japan, Mexico, Netherlands, New Zealand, Norway, Poland, Portugal, South Korea, Spain, Sweden, Switzerland, Turkey, and the United Kingdom. For more details, please visit the website of the OECD.

Our commodity price factor uses monthly changes in the Raw Industrials sub-index of the CRB Spot Commodity Index, which is available on Datastream for the period from January 1951 to December 2015.

World Equity Factor

Finally, we include a global equity factor. Whereas the correlation between equity returns and currency returns is low for developed markets (see Bekaert and Hodrick, 2011), the correlations do vary substantially through time, see for example Cappiello and De Santis (2005) and Hau and Rey (2006). We construct an equally-weighted world equity market return based on MSCI equity price indices in Australia, Canada, Europe (an index of equity markets in the Euro zone), Japan, Norway, Singapore, Sweden, Switzerland, the United Kingdom and the United States.⁴

In Table 7, we examine the explanatory power of the above existing currency factors for the variation in our 10 currency baskets. The top panel shows results for univariate regressions, the bottom panel shows the multivariate regression.

[Table 7]

In the top panel, the overwhelming majority of the factor loadings are statistically significantly different from zero for the carry, value, and equity factors. These factors have reasonably high explanatory power for the majority of the currency-basket factors. The carry and value factors explain on average 13% and 12% of the variation in the currency basket factors respectively, but the R^2 is only 4% on average for the equity factor. The average R^2 s for the other factors are even lower.

In the bottom panel, we see that the number of significant factors varies from currency to currency, being as low as 2 for the GBP and as high as 4 for the AUD and CHF. The commodity factor is surprisingly not significant for CAD. Interestingly, except for the volatility factor, every factor is significant at least once, but carry is significant for 8 and value for 9 out of 10 currencies. The R^2 for the multivariate models varies between 10% and 37% and is 23% on average.

⁴ Because equity market data for New Zealand are limited, we use equity market data for Singapore instead. We also construct a value-weighted world equity market return as the market capitalization weighted average of these equity market returns, based on equity market capital values available on Datastream. However, the equally-weighted equity market return has more explanatory power for exchange rate (co)movements. Using the MSCI world market index produces very similar results to the ones reported here.

The table reveals that the carry and value factors are the most promising candidates to feature in a factor model aimed at explaining currency comovements. However, the explanatory power is distinctly lower than the explanatory power of the new factors we proposed in Section II.

IV. The Fit of Various Factor Based Models

We have now introduced a total of 11 factors, 5 new ones and 6 factors that have been considered before, mostly in pricing exercises. Here we try to determine which model best fits the comovements across currency changes. Before we run various horse races, we examine the correlations between the factors and their relationship to the standard principal components.

IV.1. Factor Correlations

To obtain further intuition on these factors, Table 8 produces their correlation matrix and their correlations with the first three principal components of the bilateral currency changes. Note that in an N -currency world, we have $N(N-1)$ bilateral rates, but half of these are simply the negative (in logs) of another bilateral rate, leaving $N(N-1)/2$ different pairs. Thus, there are 45 ($=10 \times 9/2$) non-repeated bilateral rates among the G10 currencies. We denote the first three principal components by 45FPC, 45SPC and 45TPC, respectively.

[Table 8]

It is not surprising that our clustering technique yields a factor that is highly correlated with the first principal component (the correlation with 45FPC is -83%). However, CF_{abs} is also highly correlated with the second principal component (53%). Moreover, the currency commodity factor (CF_{com}) is more highly correlated with 45FPC (at -90%) than is our clustering factor. The CF_{TW} factor is 77% (41%) correlated with the second (third) principal component, but barely at all with the 45FPC. Recall that the trading volume-weighted factor is dominated by the dollar currency basket which implies that the first principal component in bilateral currency changes is not dominated by dollar movements. The two factors resulting from selecting three clusters, CF_3 , and CF_2 , are -37% correlated. CF_{31} includes two of important currency baskets (CB_{USD} and

CB_{JPY}), is not highly correlated with the first principal component, 45FPC, but is highly correlated with 45SPC (87%). It is CF_{32} that is very highly correlated with 45FPC (-97%)! Therefore, the cluster of AUD, CAD and NZD is the set of currencies that best approximates the first principal component in the G10 currencies.

Among the extant currency factors, the currency carry trade factor (denoted by Carry) is 62% correlated with the first principal component, 40% correlated with the second principal component, and 18% correlated with the third principal component. The currency value factor (denoted by Value) and the equally-weighted world equity market return (denoted by Equity) are respectively 56% and 33% correlated with 45FPC. The currency-volatility factor (denoted by Volatility) and the changes in the CRB Spot Commodity Index (denoted by Commodity) are respectively 25% and 33% correlated with 45SPC. The currency-momentum (denoted by Momentum) factor is not highly correlated with any of the top three principal components. It is not surprising that the carry and the value factors, which were shown to have high explanatory power for currency comovements, are relatively highly correlated with the first principal component of exchange rate movements.

To create factor models using the factors we introduced, it is important that the factors are not multi-collinear. The correlation table shows that this is clearly not the case. The highest correlations observed are those between the CF_{31} and the CF_{TW} factors (at 74%), the carry factor and the currency commodity factor (at 65%), and the currency commodity factor and the clustering factor, CF_{abs} , at 60%. There are some other high correlations (such as between CF_{com} and CF_{32}), but these factors will never be considered together.

IV.2. Horse Race between factor models

We now focus on the RMSE in correlation space to determine the factor model that best fits the currency comovements between the currency baskets. The RMSE can be viewed as the average correlation distance between the model and the data. Recall from Table 5 that the data correlations vary between -56% and +54%. Of course, the RMSE statistics are estimated from a finite data sample and we must take sampling error into account. We use the bootstrap procedure described before to determine sampling error in the RMSE's for the new factor models proposed

in this article and whether they perform significantly better than other models. Note that the sampling error of the RMSE statistic is generally relatively low at about 0.01 to 0.02, admitting relatively powerful tests.

To set the stage, Table 9 reports the RMSE for univariate factor models using all 11 factors we consider in this article. This exercise immediately reveals the value of the new CF_{abs} factor which only has a RMSE of 0.176, with the 95% confidence interval being [0.163, 0.211]. Most of the other factors have RMSE's that are far above this interval. The second best individual factor among the new factors is CF_{32} with a RMSE of 0.201. Among the extant factors, the best factor is the carry factor with a RMSE of 0.264.

In Table 10 (Panel I), we compare the fit of various multivariate models. We start with the two three factor models we proposed in Section II. The three factor models significantly reduce the RMSE, bringing it down to 0.112 for the model incorporating CF_{abs} , and to 0.131 for the model with the two clustering factors. In an absolute sense, a correlation error of about 10% seems small and these models thus match the data correlations rather well. The differences in fit between the two models seem small economically and they are also not significant in a statistical sense in that the RMSE generated by the second model is within the 95% convince interval of the first one.

The rest of this Panel investigates the fit of various combinations of the extant currency factors. When we use all 6 factors, the RMSE is 0.214, almost twice as high as the fit of our parsimonious model. When we drop the two worst performing factors (volatility and momentum), the fit does not improve. We also report the RMSE for two three-factor models adding to carry and value, either the equity factor or the commodity price factor. Both models performs similarly with an RMSE of 0.218. The RMSEs generated by these models are also outside the 95% confidence intervals generated by the bootstrap for our 2 three-factor models. We conclude that the new models we proposed are far superior to models created from extant currency factors in fitting currency comovements.

It is still conceivable that the extant currency factors can help the fit of our proposed model. We address this issue in Panel II of Table 10. We focus our attention on the carry and value factors, which are the best extant currency factors. Adding these factors does decrease the RMSE most

of the time, but the decrease is both economically and statistically insignificant, except in one case where the RMSE for the second three-factor model is 0.002 above the 95% confidence interval generated by one of the five-factor models.

We conclude that a parsimonious factor model, using a factor obtained from a simple clustering method, which groups mostly the dollar currencies, a commodity currency factor and a trading volume weighted “market” factor fits currency comovements very well and does so better than any other factor model extracted from the extant currency factors.

V. Comparison to Recent Factor Models

In this section, we first link our results to some recent academic studies regarding currency factors. We also reflect on the practicability of our factor model. Over the last few decades, increased currency trading for speculative purposes has seen the birth of several tradeable currency factors, e.g. through ETFs. We examine the link between our models and the well-known Deutsche Bank currency factors.

V.1. A Dollar Factor

The currency risk model of Lustig, Roussanov and Verdelhan (2011) introduces the U.S. dollar basket factor as a common currency factor. It is essentially the average excess return for a US investor to investing in all the foreign currencies and thus closely related to our dollar basket (CB_{USD}). Verdelhan (2016) further studies the U.S. dollar factor, showing that it accounts for a large share of bilateral exchange variations against the U.S. dollar. He identifies the U.S. dollar factor as a key risk factor and links its explanatory power for currency movements to its comovements with different macro-economic variables (in particular capital flows). Importantly, he shows that the “dollar factor” explains much more of bilateral currency comovements than does the carry factor, which we have shown to be one of the better extant currency factors.

In this section, we show that the dollar factor’s explanatory power measured in Verdelhan (2016) is potentially spurious. In addition, Verdelhan’s dollar factor is numeraire dependent and

therefore fits the currency movements in other countries poorly. Verdelhan does create a factor that better captures global currency movements but it is still closely related to the dollar basket. We show that this factor is dominated by our CF_{abs} factor.

Consider the main regression in Verdelhan (2016):

$$\Delta s_{k,USD,t+1} = \alpha + \beta Carry_{t+1} + \gamma CB_{USD,t+1} + \epsilon_{t+1} \quad (23)$$

where, $Carry_{t+1}$ is the difference in returns between portfolios of high and low interest rate currencies. Now, recall Equation (7) holding that each bilateral exchange rate change can be described as (9/10 times) the difference between the two corresponding currency baskets. Therefore:

$$\frac{9}{10} CB_{USD,t+1} - \frac{9}{10} CB_{k,t+1} = \alpha + \beta Carry_{t+1} + \gamma CB_{USD,t+1} + \epsilon_{t+1}, \quad (24)$$

or:

$$CB_{k,t+1} = -\frac{10}{9}\alpha - \frac{10}{9}\beta Carry_{t+1} + (1 - \frac{10}{9}\gamma) CB_{USD,t+1} - \frac{10}{9}\epsilon_{t+1}. \quad (25)$$

The results of regressions (24) and (26) are reported respectively in Panels I and II of Table 11. As can be seen, the coefficients of carry in the right panel are almost equal to “ $-\frac{10}{9}$,” times the coefficients of carry in the left panel. In addition, the coefficients of CB_{USD} in the right panel are almost equal to one plus “ $-\frac{10}{9}$,” times the coefficients of CB_{USD} in the left panel. Finally, the adjusted R-squares in the left panel are much higher than those in the right panel. The presence of a common component on the left hand side and the right hand side in Equation (23) leads to a somewhat different interpretation of the Verdelhan – results. First, the coefficients in Regression (23) are difficult to interpret. For example, the “dollar factor” has virtually no independent effect on CB_{GBP} yielding an insignificant -0.06 coefficient; yet, regression (23) produces a coefficient of 0.95 (which is, of course nothing but 9/10 minus 9/10 times -0.06), with a huge t-statistic. In addition, the carry coefficient is difficult to interpret, because it implicitly really reflects the (negative of) the dependence between the appreciation of currency k w.r.t a basket of currencies and carry. Second, the explanatory power of the carry and the CB_{USD} for bilateral exchange rates, using Equation (23) (in the left panel) is artificially high because we use a component in the left

hand-side variable as a right-hand-side explanatory variable. For example, in the last row of Panels I and II, the R^2 of carry and CB_{USD} is 57% for the changes in the GBP/USD while in fact the R^2 of carry and CB_{USD} is only 1% for CB_{GBP} . In contrast, in the second row, the R^2 of carry and CB_{USD} is the lowest at 39% for the changes in the CAD/USD while in fact the R^2 of carry and CB_{USD} is the second highest at 41% for the CB_{CAD} .

Perhaps aware of the potential problem, Verdelhan (2016) excludes the left-hand-side exchange rate in the composition of his dollar factor, but it is easy to see that this does not resolve the problem.⁵

[Table 11]

Moreover, this now aggravates the problem that the factor is not common across even bilateral rates relative to the dollar. Furthermore, as we have shown before, the original dollar factor, CB_{USD} , is not a suitable common factor for all bilateral rates.

To resolve the numeraire currency problem, Verdelhan (2016) proposes two alternatives. First, use the numeraire currency basket factor as the explanatory variable; for example a pound basket factor for the bilateral rates w.r.t. the British pound. Obviously, such factors are not truly global and all will perform poorly in terms of global fit (see Table 3). Second, in an attempt to find a really global factor, Verdelhan (2016) proposes the difference in exchange rate changes between high and low dollar beta portfolios, hereafter denoted by HML Dollar or HML\$. To create such a portfolio, he regresses currency changes in a rolling fashion on the carry and dollar baskets and sorts currencies according to their dollar basket exposures. From our analysis above, this exercise essentially sorts on the dollar basket exposure of other currency baskets and is therefore potentially a valid global risk factor.

⁵ Assume that EUR/USD is the left-hand-side variable in the regression. If we exclude it in the composition of the dollar basket factor as well as in the euro basket factor, we have:

$CB_{\$} = \frac{1}{8} [\sum_{j \neq \epsilon}^8 \Delta s_{j,\$}]$ and $CB_{\epsilon} = \frac{1}{8} [\sum_{j \neq \$}^8 \Delta s_{\epsilon,j}]$, and as a result, $\Delta s_{\epsilon,\$} = \frac{8}{9} CB_{\$} - \frac{8}{9} CB_{\epsilon}$. As can be seen, the dollar basket factor ($CB_{\$}$) is still a part of left-hand-side variable. Thus, our concern is valid even after excluding the EUR/USD exchange rate changes ($\Delta s_{\epsilon,\$}$ and $\Delta s_{\$, \epsilon}$) in the composition of the basket factors, $CB_{\$}$ and CB_{ϵ} respectively. The supportive empirical evidence is available on request.

However, we now show that our simple currency factor (CF_{abs}) has more explanatory power for currency variation than the $HML_{\$}$ factor of Verdelhan (2016). We run the following horserace regressions:

$$CB_{k,t+1} = \alpha + \beta HML_{\$t+1} + \gamma CF_{abs,t+1} + \epsilon_{t+1}. \quad (26)$$

The results are reported in Table 12. Panel I and II show the explanatory power of $HML_{\$}$ and CF_{abs} separately and Panel III shows the explanatory power of $HML_{\$}$ and CF_{abs} jointly. The adjusted R^2 reported in panel II are mostly higher than in Panel I. The $HML_{\$}$ has, not surprisingly, much higher explanatory power (R^2 of 71%) for the US dollar basket factor (CB_{USD}) while CF_{abs} has more balanced explanatory power for most of the rates compared to the $HML_{\$}$ factor (an average R^2 of 33% versus only 23% for $HML_{\$}$). Furthermore, the coefficient of $HML_{\$}$ is significant for only 7 out of 10 currency baskets while the coefficient of our CF_{abs} factor is significant for all G10 currency baskets.

When we put both factors together in Panel III, the adjusted R^2 s increase, which shows that the two factors contain different information. The coefficient of CF_{abs} remains statistically significant for all currency baskets while the coefficient of $HML_{\$}$ is not significant for the CHF basket factor (CB_{CHF}).

The last row of the table reports the results of the comovement fitting horse race. The $HML_{\$}$ factor has a RMSE of 0.304 relative to the data correlations, which is much higher than the 0.179 RMSE generated by our CF_{abs} factor. Moreover, the bivariate model has a worse RMSE (of 0.266) than the CF factor by itself. We conclude that the explanatory power of our simple currency factor (CF_{abs}) is higher than that of the global dollar factor of Verdelhan (2016) and there is no reason to consider the $HML_{\$}$ factor for a successful factor model.

We also examined the explanatory power of dollar-carry factor introduced in Lustig et al. (2014), which goes long in a basket of foreign currencies and short in the dollar whenever the average foreign short-term interest rate is above the U.S. interest rate and vice versa. However, the explanatory power of the dollar-factor is much lower than even the $HML_{\$}$. The results are reported in the online Appendix.

[Table 12]

V.2. Re-interpreting the currency factor structure in Lustig and Richmond (2016)

Lustig and Richmond (2016) recently detect an interesting pattern in cross-currency correlations. They regress bilateral exchange rate changes on “base factors,” which are closely related to our currency baskets. They then show that the betas in these regressions and the R^2 s can be interpreted using a gravity model: they are lower the closer the countries are in terms of distance and other variables measuring economic closeness.

At first glance, this is somewhat surprising. For example, it is well – known that the variability of exchange rate changes between closely connected countries (for example, Canada and the US; countries within Europe) display less variability which is typically interpreted as reflecting close economic ties. Rose, 2000, and Baxter and Kouparitsas (2006), for example, show that FX volatility is negatively correlated with trade volume. In the Lustig-Richmond world, the currencies of these countries show low betas and low systematic risk. It is also surprising perhaps from the perspective of our clustering model, where we measure distance as a negative function of correlation and find a block structure that puts countries that are geographically close (the European currencies) within one block (at least when we use two clusters).

However, our results can be reconciled. To illustrate this, let’s consider the US dollar as the base currency. In that case, the base factor is in fact equivalent to our USD currency basket. The regression that Lustig and Richmond run is essentially a regression of bilateral currency changes relative to the dollar onto the dollar currency basket. We replicate such a regression for the 9 other currencies among the G10 in Table 13, Panel I. The results, at least informally, confirm the Lustig-Richmond results. Within Europe, the beta is smallest for the economically close UK and higher for the further away Scandinavian countries. The beta is by far the lowest for the neighboring country. In this article, we opted to think about currency baskets globally to represent currency movements. From that perspective, the USD and CAD currency baskets are positively correlated (they show a short “distance”) as are the JPY and USD baskets (see Table 3). The JPY and CAD have the lowest betas and R^2 s in the Lustig -Richmond regressions. Panel II shows that if we recast the Lustig-Richmond regressions in our currency basket framework, the CAD and JPY baskets are the only baskets with positive betas. Thus, our results can in fact be

reconciled and the Lustig-Richmond framework provides economic intuition for our correlation structure.⁶

To further interpret the regressions in Panel I, note that the left-hand side variable is also contained in the independent variable. This would naturally bias the betas downward for neighboring countries because of their lower variability. Lustig and Richmond (2016) control for this bias by excluding the bilateral exchange rate on the left from the right-hand side base factor. Nevertheless, this does not quite fully control for the variability bias. Exchange rate changes that show low absolute variability will necessarily also show lower covariances with other variables (and the variance of the independent variable is only marginally affected by the identity of the dependent variable).

Mechanically, our currency basket concept also sheds light on the Lustig-Richmond findings. Intuitively, currency baskets of nearby countries should be highly correlated (and they are, see Table 5). How can the regression beta in the Lustig-Richmond regression then be positively correlated with distance? Well, recall that a bilateral currency change is approximately the difference between two currency baskets. Therefore, the Lustig-Richmond regression regresses the difference between the dollar basket and another currency basket onto essentially the dollar basket. Therefore, the regression beta will be decreasing in the comovement between the currency baskets.

Thus, our distance concept applied to currency baskets is intuitively positively correlated with physical distance. For example, when we run a regression of the bilateral distances (BDistance) between the currency baskets as used in the clustering algorithm on the population weighted physical distance (PDistance) between the involved countries (as used in Lustig and Richmond), we obtain:

$$\text{BDistance}_{i,j} = 1.33 + 1.86 \times 10^{-5} \text{ PDistance}_{i,j} + e \quad (R^2=0.09)$$

t-statistics: (95.56) (7.77)

⁶ Examining other currency perspectives (such as the Euro) revealed that the Lustig-Richmond intuition does not uniformly survive. In contrast, our factor structure, with currency baskets in Europe positively correlated within and negatively correlated outside Europe, is a global feature of currencies.

This means that for every physical mile the “correlation” distance increases by 1.86×10^{-5} ; the coefficient is highly statistically significant. We also run the same regression using actual correlations between the currency baskets finding statistically significant negative coefficients.

V.3. On The Deutsche Bank Factors

In the “active” currency management space, the Deutsche Bank (DB) currency factors are frequently used to measure risk exposures and performance. Levich and Pojarliev (2008) show a negative correlation between the performance of currency managers and the R^2 of a regression of their returns onto the DB factors. The DB factors include a carry, momentum and value factor. The Appendix briefly discuss the construction of these factors, which is slightly different from our construction of the analogous factors. For example, the DB carry factor is constructed by ranking currencies on 3-month Libor rates, while we use government bond rates that are available for a longer period of time. The DB momentum factor is constructed by ranking currencies according to their 12 month rolling return against the USD, while we rank currencies according to their 1 month rolling returns. Yet, the factors show relatively high correlation. Over the sample period available,⁷ the carry factors are 87% correlated, the value factors 88% correlated but the momentum factors are only 20% correlated.

Because of the relatively high correlation, it is to be expected that the DB factors perform about the same as the analogous standard currency factors that we examined before. Online Appendix Table 6 confirms that a three-factor model using the DB factors performs similarly relative to the analogous model using our factors, in terms of R^2 and is slightly worse in terms of the RMSE for correlations.

This also implies that the performance of the DB factors falls far short of the performance of our new three-factor model involving the clustering, commodity and market factors. It would therefore be of interest to create tradeable baskets corresponding to our three factors. We discussed before that our “cluster factor” involves shorting (longing) the European (dollar) currencies. Likewise, a proxy to the trade-weighted world factor can be obtained by going long in

⁷ The DB FX factors are available for a shorter sample period. For example, the DB’s momentum and value factors are available since July 1989 and the DB’s carry factor is available since October 2000.

all foreign currencies relative to the USD, roughly proportional to their trade weights (see Appendix B).

VI. Explaining Emerging Market Currencies

As an out-of-sample exercise, we verify how well our new factor models and the extant currency factors fit exchange rate variation and currency components in 14 emerging markets. Our sample period for the emerging markets extends from January 1990 to December 2015.⁸ We consider our two new models, one involving CF_{abs} , the currency commodity factor (CF_{com}) and the market factor (CF_{tw}); the other involving the two cluster factors (CF_{31} and CF_{32}) and the market factor. We also consider the performance of our cluster factors separately.

Our set of emerging countries includes China, Czech Republic, Hungary, India, Indonesia, Mexico, Philippines, Poland, Singapore, South Africa, South Korea, Taiwan, Thailand, and Turkey. Note that our currency baskets in this case include both emerging and developed currencies (that is, 24 currencies). For example, the China Yuan basket factor is the average appreciation of Chinese yuan with respect to 13 other emerging currencies and the 10 G10 currencies. The correlations we consider however, are the correlations between the emerging market baskets.

Table 14 (Panel I) presents the results of univariate regressions of the 14 emerging currency-baskets on CF_{abs} . A single CF_{abs} factor model explains on average 6% of the variation in emerging currency basket changes and the coefficient on CF_{abs} is statistically significant for 12 out of 14 emerging currency baskets. The RMSE of its implied currency correlations is 0.159. Therefore, the CF_{abs} factor has less power to explain variation in emerging currency basket changes than in G10 currency baskets. However, it fits the correlations among the 14 emerging currency baskets slightly better than those among G10 currency baskets. While surprising at first, the online Appendix shows that the emerging market baskets show relatively small correlations due to the large idiosyncratic risk displayed by emerging market currencies. This fact both explains the low R^2 's and makes the correlation error larger in relative terms.

⁸ Because the sample is shorter than for the G10 currencies, the results are not exactly comparable with our previous results. However, in results available upon request, we find that the RMSE's for the shorter sample are similar (albeit slightly higher than) that for the full sample (0.12-0.13 for the two three-factor models).

[Table 14]

Panel II presents the results of bivariate regressions of the 14 emerging currency-basket changes on the two factors from the three-absolute clusters, CF_{31} and CF_{32} . The two factor model explains on average 13% of currency variation and the coefficients for both factors are statistically significant for 11 out of 14 emerging currency baskets. The RMSE of its implied currency correlations is 0.121. We ran a block bootstrap for this set of currencies too. Using the 95% confidence intervals for this model, the drop from 0.159 to 0.121 is actually statistically significant.

We now consider the performance of the two new three factor models. Panel III presents the results of multivariate regressions of the 14 emerging currency-baskets on our suggested three factors CF_{abs} , CF_{com} , and CF_{TW} jointly. The commodity factor, CF_{com} , is statistically significant for 6 currencies and the CF_{TW} is significant for 7 currencies. Our new three-factor model explains on average 15% of the emerging currency basket return variation and the RMSE of its implied currency correlations is 0.121.

In Panel IV, we use the other three factor model, combining our currency trading-volume weighted factor, CF_{TW} , with CF_{31} and CF_{32} . This model generates an average R^2 of 16% and the coefficient on CF_{TW} is statistically significant for 6 out of 14 emerging currency baskets. The RMSE of the model's implied currency correlations is 0.121, which is only slightly lower than that of the two factor model. Using our bootstrap exercise, the improvements produced by the two three factor models relative to the model with the two cluster factors are economically and statistically insignificant.

It is undoubtedly true that emerging market currencies are quite different from developed market currencies, featuring lower trading volumes, higher spreads, and more often than not more government meddling. Moreover, many of these countries experience larger economic shocks than developed countries do and emerging market currency values may reflect the economic prospects of a country more than developed market currencies do. Interestingly, it is well known that equity and currency markets are more correlated in emerging than in developed markets, for example (see Bekaert and Harvey, 2015). All these factors may cause emerging market currencies to have more country specific risk. However, it is also conceivable that there is an

emerging market factor. Emerging markets is still a popular asset class among institutional investors, and currency hedging for emerging market investments remains uncommon, so that large movements in the asset class may be reflected in emerging market currencies overall. Moreover, these currencies also tend to have properties of a special “carry trade,” with many emerging market currencies featuring high interest rates.

We therefore create an emerging market currency factor (CF_{EM}) using relative trade-weights as we did for the G10 currencies. The trade weights are reported in the appendix. In the latest BIS survey the Mexican peso was the most traded emerging market currency (almost 19% of trading volume), followed by the Chinese Renminbi (16.5%) and the Singapore dollar (10.5%). Earlier in the sample, the SGD was the dominant currency, but the Mexican peso always has comprised an important part of trading volumes. In contrast, the CNY represented a rather negligible part of trading volumes up and till 2007 but since then its weight has increased to over 15%.

In Panel V, we replace the CF_{TW} factor in our preferred three factor model by the corresponding emerging market factor and in Panel VI we add it to the basic three factor model. However, replacing the CF_{TW} factor by the emerging market factor worsens the performance of the model, with the R^2 failing to improve and the RMSE becoming slightly worse. When we use the 4-factor model, the R^2 modestly increases to 18% but the correlation fit is still worse than for the parsimonious three factor model. Yet, 10 of the emerging market currencies have significant exposure to the emerging market factor! We performed the same analysis for the alternative three factor model with CF_{31} and CF_{32} . The results in Panel VII and VIII are very similar to those in Panels V and VI respectively, but the four factor model now also does marginally better in terms of RMSE.

VII. Conclusions

We examine various factor models to explain currency (co) movements and document their fit with the data. Rather than studying bilateral rates with a specific numeraire currency, as is customary in the literature, we focus on “currency baskets,” which can be viewed as the principal components for each currency’s changes relative to all other currencies. For the G-10 currencies, studying 10 currency baskets is equivalent to studying all 45 unique currency pairs. This

methodology together with a clustering technique helps us detect a clear factor structure in currency comovements suggesting two currency blocks. One block includes the dollar rates and the yen the other block includes other major currencies, in particular European currencies.

The new factor is a very significant determinant of variation in the 10 currency baskets. When combined with a currency commodity factor (including the AUD, NZD, CAD and the NOK) and a market factor, which we construct from the currency baskets using trade volumes from the BIS, a parsimonious factor model results that explains on average 58% of the changes in the various currency baskets. It also fits the currency correlations quite well, generating a RMSE of only 0.11. In addition, this parsimonious model also has significant explanatory power for emerging currency baskets and fits their comovements well, with a RMSE of only 0.12.

We also compare the performance of the new model with that of extant currency factors. The extant currency factors in our tests are the carry, volatility, value, and momentum currency factors, a commodity prices factor, and a world-equity factor. We contrast the explanatory power of the extant currency factors with the explanatory power of various principal components. The carry and value factors fit currency variances and correlation much better than the other currency extant factors. However, any factor model created from the extant currency factors performs much less well than the new factor model. This result also extends to emerging market baskets.

We also find that our new currency factor—the sum of dollars basket factors— has more explanatory power compared to the global dollar factor of Verdelhan (2016).

With active currency management becoming more commonplace, our findings should help currency managers and international investors to better explain the risks and comovements of currencies worldwide. The new factor model can also be help in international asset pricing exercises, where a multitude of bilateral currency rates can be collapsed into the factors we uncovered. Of course, much additional research is needed. We have only studied our factors in terms of their ability to fit comovements, and have not consider the returns associate with them or their ability to price the cross-section of currency portfolios. Finally, we have also focused on unconditional correlations, and it is well known that currency correlations vary through time (see e.g. Hau and Rey, 2006). It should be straightforward to use high frequency data to extend our methodology to a conditional framework.

References

- Abuaf, N. and Jorion, P. Purchasing Power Parity in the Long Run. *Journal of Finance* 45, 1 (1990), 157–174.
- Adler, M. and B. Dumas. International portfolio choice and corporation Finance: A synthesis. *The Journal of Finance* 38, 3 (1983), pp. 925–984.
- Ahn, D., J. Conrad, and R. F. Dittmar, Basis Assets. *Review of Financial Studies* 22, 12 (2009), 5133–5174.
- Aloosh, A. Global Variance Risk Premium and Forex Return Predictability. *Available at SSRN 2133999*, (2012).
- Bakshi, G. and Panayotov, G. Predictability of Currency Carry Trades and Asset Pricing Implications. *Journal of Financial Economics* 110, 1 (2013), 139–163.
- Baxter, Marianne and Michael Kouparitsas. What determines bilateral trade flows?, NBER Working Paper 12188, (2006).
- Bekaert, Geert, and Campbell R. Harvey. Emerging Equity Markets in a Globalizing World. *Available at SSRN 2344817*, (2015).
- Bekaert, Geert, and Robert J. Hodrick. International Financial Management, Prentice Hall, 2nd edition, (2011).
- Bekaert, G., R. J. Hodrick, and X. Zhang. International Stock Return Comovements. *Journal of Finance* 64 (2009), 2591–2626.
- Brusa, F., T. Ramadorai, and A. Verdelhan. The International CAPM redux, *Available at SSRN 2462843*, (2014).
- Burnside, C., Eichenbaum, M., Rebelo, S. Carry trade and momentum in currency markets. *Annual Review of Financial Economics* 3, (2011), 511–535.
- Cappiello, L., and De Santis, R. A. Explaining exchange rate dynamics: The uncovered equity return parity condition. *SSRN eLibrary* (2005).
- Chen, Y., and Rogoff, K. Commodity currencies. *Journal of International Economics* 60, (2003), 133–160.
- Cochrane, J. H. Asset Pricing. *Princeton, NJ: Princeton University Press*, (2001).

- Connor, G. and Korajczyk, R. Performance Measurement with the Arbitrage Pricing Theory: A New Framework for Analysis. *Journal of Financial Economics* 15, 3 (1986), 373–394.
- Connor, G. and Korajczyk, R. Risk and Return in an Equilibrium APT: Application of a New Test Methodology. *Journal of Financial Economics* 21, 2 (1988), 255–289.
- Connor, G. and Korajczyk, R. A Test for the Number of Factors in an Approximate Factor Model. *Journal of Finance* 48, 4 (1993), 1263–1291.
- Daniel, K., and Titman, S. "Testing Factor-Model Explanations of Market Anomalies", *Critical Finance Review* 1, 1 (2012), 103–139
- Engel, C., Mark, N. C., and West, K. D. Factor Model Forecasts of Exchange Rates. *Econometric Reviews* 34, 1-2 (2015), 32–55.
- Fama, E., and French, K. Multifactor Explanations of Asset Pricing Anomalies. *Journal of Finance*, 51 (1996), 55–84.
- Fama, E. F., and K. R. French. Value versus Growth: The International Evidence. *Journal of Finance* 53, (1998), 1975–1999.
- Hafeez, B., 2007. Deutsche Bank Guide to Currency Indices, Section Therein Titled “Currency Markets: Money Left on the Table?” Deutsche Bank.
- Hau, H., and Rey, H. Exchange rates, equity prices, and capital flows. *Review of Financial Studies* 19, 1 (2006), 273–317.
- Hou, K., Karolyi, G.A., Kho, B.C. What factors drive global stock returns? *Review of Financial Studies* 24, 3 (2011), 2527–2574.
- Jagannathan, R., and Wang, Z. The Conditional CAPM and the Cross-Section of Expected Returns. *Journal of Finance*, 51 (1996), 3–54.
- Levich, R., and Pojarliev, M. Do professional currency managers beat the benchmark? *Financial Analysts Journal* 65, (2008), 18–32.
- Lo, A., and MacKinlay, A.C. Data-Snooping Biases in Tests of Financial Asset Pricing Models. *Review of Financial Studies*, 3 (1990), 431–68.
- Lothian, J. R., and Taylor. M. P. Real Exchange Rate Behavior: The Recent Float from the Perspective of the Past Two Centuries. *Journal of Political Economy* 104, (1996), 488–509.
- Lustig, H., and Richmond, R. Gravity in FX Rsquared: Understanding the Factor Structure in Exchange Rates. Working paper, (2015).

- Lustig, H., Roussanov, N., and Verdelhan, A. Countercyclical currency risk premia. *Journal of Financial Economics* 111, 3 (2014), 527–553.
- Lustig, H., Roussanov, N., and Verdelhan, A. Common risk factors in currency markets. *The Review of Financial Studies* 24, 11 (2011), 3731–3777.
- Mark, N., Exchange Rates and Fundamentals: Evidence on Long-Horizon Predictability, *American Economic Review* 85, (1995), 201–218.
- Markowitz, H. Portfolio Selection. *Journal of Finance*, 7 (1952), 77–91.
- Menkhoff, L., Sarno, L., Schmeling, M., and Schrimpf, A. Carry trades and global foreign exchange volatility. *The Journal of Finance* 67, 2 (2012a), 681–718.
- Menkhoff, L., Sarno, L., Schmeling, M., and Schrimpf, A. Currency momentum strategies. *Journal of Financial Economics* 106, 3 (2012b), 660–684.
- Menkhoff, L., Sarno, L., Schmeling, M., and Schrimpf, A. Currency Value. *The Review of Financial Studies* 30, 2 (2017), 416–441.
- Ormerod, P., and Mounfield, C. Localised Structures in the Temporal Evolution of Asset Prices. *In New Approaches to Financial Economics*, Santa Fe Conference, (2000).
- Ready, R., Roussanov, N., and Ward, C. Commodity Trade and the Carry Trade: a Tale of Two Countries. NBER Working Paper 19371, (2013).
- Rime, D., and Schrimpf A. The anatomy of the global FX market through the lens of the 2013 Triennial Survey. *BIS Quarterly Review*, (2013).
- Rose, Andrew K. One money, one market: estimating the effect of common currency on trade, *Economic Policy*, vol. 30 (2000), 7–45.
- Ross, S. Mutual Fund Separation in Financial Theory: The Separating Distributions. *Journal of Economic Theory*, 17 (1978), 254–86.
- Stock, J.H., and Watson, M.W. Diffusion Indexes. NBER Working Paper 6702 (1998).
- Verdelhan, A. The share of systematic variation in bilateral exchange rates. *The Journal of Finance*, forthcoming (2016).

Appendix A

Tables A1 and A2 present the currency trading weights reported by the BIS (Bank for International Settlements). Table A3 reports the explanatory power of the trade-weighted currency factor (CF_{TW}). As the USD is the most traded currency (more than 47% of all currency trading volumes), it has a high weight in the construction of CF_{TW} and thus the CF_{TW} has a high explanatory power for CB_{USD} , with an R^2 of 83%. However, CF_{TW} has lower average explanatory power than CF_{abs} (see Table 6).

Table A1 – The G10 Currency Trade Weights

Date	USD	AUD	CAD	CHF	EUR	JPY	NOK	SEK	NZD	GBP
30.04.1998	50.6%	1.7%	2.0%	4.1%	22.1%	12.6%	0.1%	0.2%	0.1%	6.4%
30.04.2001	48.9%	2.3%	2.4%	3.3%	20.6%	12.8%	0.8%	1.4%	0.3%	7.1%
30.04.2004	47.9%	3.3%	2.3%	3.3%	20.4%	11.3%	0.8%	1.2%	0.6%	9.0%
30.04.2007	47.8%	3.7%	2.4%	3.8%	20.7%	9.6%	1.2%	1.5%	1.1%	8.3%
30.04.2010	47.1%	4.2%	2.9%	3.5%	21.7%	10.5%	0.7%	1.2%	0.9%	7.2%
30.04.2013	48.7%	4.8%	2.6%	2.9%	18.7%	12.9%	0.8%	1.0%	1.1%	6.6%

This table report currency trade-weights for the G10 currencies. The weights are based on trade volumes reported by BIS (Bank for International Settlements). The trade weights are re-normalized to add to 1.

Table A2 – Emerging Currency Trade Weights

Date	CNY	CZK	HUF	INR	IDR	MXN	PHP	PLN	SGD	ZAR	KRW	TWD	THB	TRY
30.04.1998	0.0%	10.0%	0.0%	3.3%	3.3%	16.7%	0.0%	3.3%	36.7%	13.3%	6.7%	3.3%	3.3%	0.0%
30.04.2001	0.0%	4.0%	0.0%	4.0%	0.0%	16.0%	0.0%	10.0%	22.0%	18.0%	16.0%	6.0%	4.0%	0.0%
30.04.2004	1.7%	3.4%	3.4%	5.2%	1.7%	19.0%	0.0%	6.9%	15.5%	12.1%	19.0%	6.9%	3.4%	1.7%
30.04.2007	6.2%	2.5%	3.7%	8.6%	1.2%	16.0%	1.2%	9.9%	14.8%	11.1%	14.8%	4.9%	2.5%	2.5%
30.04.2010	9.0%	2.0%	4.0%	10.0%	2.0%	13.0%	2.0%	8.0%	14.0%	7.0%	15.0%	5.0%	2.0%	7.0%
30.04.2013	16.5%	3.0%	3.0%	7.5%	1.5%	18.8%	0.8%	5.3%	10.5%	8.3%	9.0%	3.8%	2.3%	9.8%

This table report currency trade-weights for emerging currencies. The weights are based on trade volumes reported by BIS (Bank for International Settlements). The trade weights are re-normalized to add to 1.

Table A3 - Explanatory Power of the Trade-Weighted Currency Factor (CF_{TW}) for G10 Currency Baskets.

	CF_{TW}	R^2
CB_{USD}	1.95 -45.38	0.83
CB_{AUD}	-0.74 (-4.92)	0.08
CB_{CAD}	0.68 -5.26	0.11
CB_{CHF}	-0.4 (-3.07)	0.03
CB_{EUR}	-0.18 (-2.43)	0.01
CB_{JPY}	0.98 -5.75	0.12
CB_{NOK}	-0.61 (-6.79)	0.12
CB_{SEK}	-0.68 (-5.69)	0.13
CB_{NZD}	-1.05 (-6.05)	0.16
CB_{GBP}	0.04 -0.35	0.00
Equally-Weighted R^2		0.16
Trade-Weighted R^2		0.43

The table presents results of regressing currency basket returns on our trade-weighted currency factors (CF_{TW}). The t-statistics are based on White standard errors and reported in parentheses. The table also reports both the equally-weighted and trade-weighted average adjusted R^2 . The sample extends from 01/1973 to 12/2015.

Appendix B

It is easy to construct tradeable indices for our factors. For example for US investors we have:

$$\begin{aligned} CB_{com} &= 4 (CB_{USD}) - \frac{10}{9} (\Delta s_{AUD,USD} + \Delta s_{CAD,USD} + \Delta s_{NOK,USD} + \Delta s_{NZD,USD}), \\ CB_{com} &= \frac{4}{9} (\Delta s_{EUR,USD} + \Delta s_{JPY,USD} + \Delta s_{SEK,USD} + \Delta s_{CHF,USD} + \Delta s_{GBP,USD}) \\ &\quad - \frac{6}{9} (\Delta s_{AUD,USD} + \Delta s_{CAD,USD} + \Delta s_{NOK,USD} + \Delta s_{NZD,USD}). \end{aligned} \quad (A.1)$$

$$\begin{aligned} CB_{31} &= 3 (CB_{USD}) - \frac{10}{9} (\Delta s_{GBP,USD} + \Delta s_{JPY,USD}), \\ CB_{31} &= \frac{3}{9} (\Delta s_{EUR,USD} + \Delta s_{SEK,USD} + \Delta s_{CHF,USD} + \Delta s_{AUD,USD} + \Delta s_{CAD,USD} + \Delta s_{NOK,USD} + \Delta s_{NZD,USD}) \\ &\quad - \frac{7}{9} (\Delta s_{GBP,USD} + \Delta s_{JPY,USD}). \end{aligned} \quad (A.2)$$

$$\begin{aligned} CB_{32} &= 3 (CB_{USD}) - \frac{10}{9} (\Delta s_{AUD,USD} + \Delta s_{CAD,USD} + \Delta s_{NZD,USD}), \\ CB_{32} &= \frac{3}{9} (\Delta s_{EUR,USD} + \Delta s_{JPY,USD} + \Delta s_{SEK,USD} + \Delta s_{CHF,USD} + \Delta s_{GBP,USD} + \Delta s_{NOK,USD}) \\ &\quad - \frac{7}{9} (\Delta s_{AUD,USD} + \Delta s_{CAD,USD} + \Delta s_{NZD,USD}). \end{aligned} \quad (A.3)$$

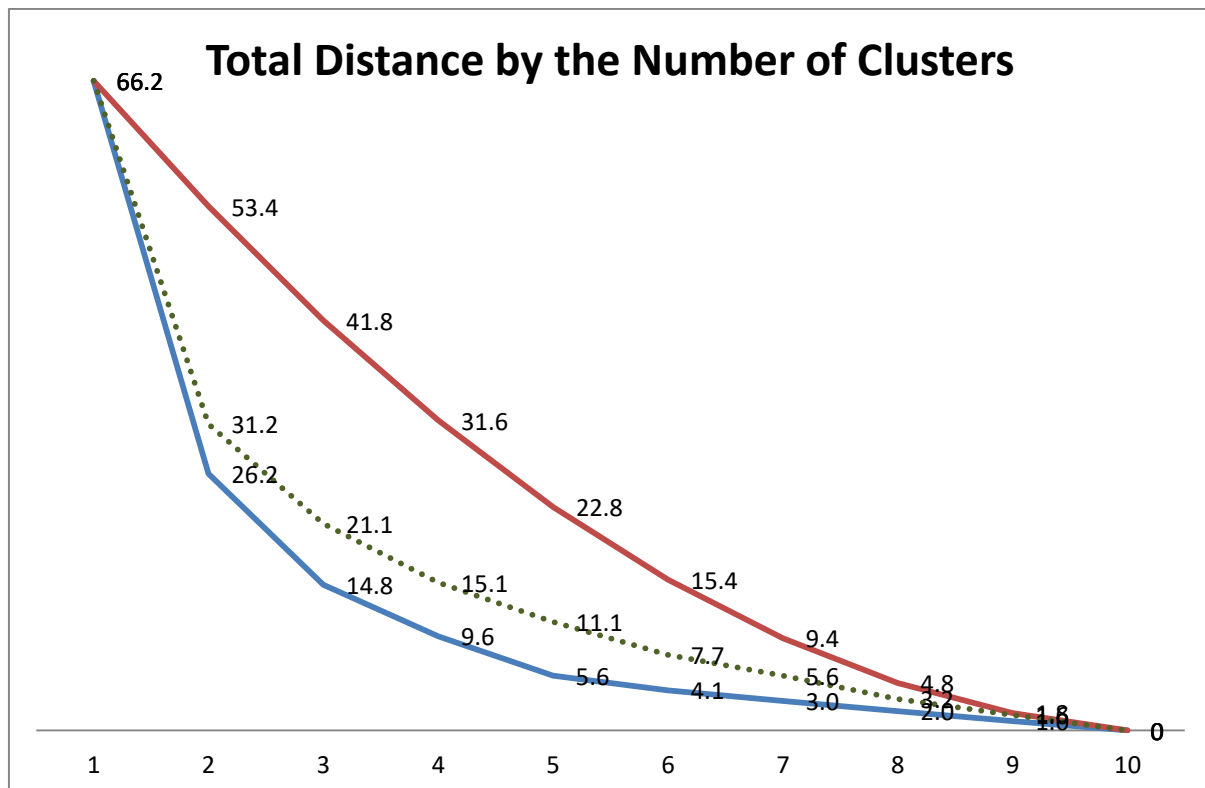
$$\begin{aligned} CB_{TW} &= (w_{USD} + w_{AUD} + w_{CAD} + w_{CHF} + w_{EUR} + w_{JPY} + w_{NOK} + w_{SEK} + w_{NZD} + w_{GBP}) (CB_{USD}) \\ &\quad - \frac{10 w_{AUD}}{9} \Delta s_{AUD,USD} - \frac{10 w_{CAD}}{9} \Delta s_{CAD,USD} - \frac{10 w_{CHF}}{9} \Delta s_{CHF,USD} - \frac{10 w_{EUR}}{9} \Delta s_{EUR,USD} \\ &\quad - \frac{10 w_{JPY}}{9} \Delta s_{JPY,USD} - \frac{10 w_{NOK}}{9} \Delta s_{NOK,USD} - \frac{10 w_{SEK}}{9} \Delta s_{SEK,USD} - \frac{10 w_{NZD}}{9} \Delta s_{NZD,USD} \\ &\quad - \frac{10 w_{GBP}}{9} \Delta s_{GBP,USD}, \end{aligned}$$

as $w_{USD} + w_{AUD} + w_{CAD} + w_{CHF} + w_{EUR} + w_{JPY} + w_{NOK} + w_{SEK} + w_{NZD} + w_{GBP} = 1$, we get:

$$\begin{aligned} CB_{TW} &= \frac{1-10 w_{AUD}}{9} \Delta s_{AUD,USD} + \frac{1-10 w_{CAD}}{9} \Delta s_{CAD,USD} + \frac{1-10 w_{CHF}}{9} \Delta s_{CHF,USD} \\ &\quad + \frac{1-10 w_{EUR}}{9} \Delta s_{EUR,USD} + \frac{1-10 w_{JPY}}{9} \Delta s_{JPY,USD} + \frac{1-10 w_{NOK}}{9} \Delta s_{NOK,USD} + \frac{1-10 w_{SEK}}{9} \Delta s_{SEK,USD} \\ &\quad + \frac{1-10 w_{NZD}}{9} \Delta s_{NZD,USD} + \frac{1-10 w_{GBP}}{9} \Delta s_{GBP,USD}. \end{aligned} \quad (A.4)$$

The weights can be rescaled to represent unit dollar investments.

Figure 1 – How Optimal Clustering Lowers Distance



The figure presents the minimum (blue line), median (dotted green line), and maximum (red line) distance for various numbers of absolute clusters of G10 currency baskets. The distances are computed as in Equation (13).

Table 1 - Bilateral Currency Correlations.

Panel I	$\Delta S_{AUD,USD}$	$\Delta S_{CAD,USD}$	$\Delta S_{EUR,USD}$	$\Delta S_{JPY,USD}$	$\Delta S_{NZD,USD}$	$\Delta S_{NOK,USD}$	$\Delta S_{SEK,USD}$	$\Delta S_{CHF,USD}$
$\Delta S_{CAD,USD}$	0.55							
$\Delta S_{EUR,USD}$	0.40	0.33						
$\Delta S_{JPY,USD}$	0.17	0.03	0.45					
$\Delta S_{NZD,USD}$	0.69	0.43	0.47	0.25				
$\Delta S_{NOK,USD}$	0.44	0.41	0.83	0.38	0.46			
$\Delta S_{SEK,USD}$	0.43	0.40	0.81	0.37	0.47	0.82		
$\Delta S_{CHF,USD}$	0.31	0.20	0.84	0.53	0.41	0.74	0.71	
$\Delta S_{GBP,USD}$	0.34	0.31	0.69	0.36	0.42	0.66	0.63	0.60
Panel II	$\Delta S_{AUD,JPY}$	$\Delta S_{CAD,JPY}$	$\Delta S_{EUR,JPY}$	$\Delta S_{USD,JPY}$	$\Delta S_{NZD,JPY}$	$\Delta S_{NOK,JPY}$	$\Delta S_{SEK,JPY}$	$\Delta S_{CHF,JPY}$
$\Delta S_{CAD,JPY}$	0.77							
$\Delta S_{EUR,JPY}$	0.60	0.65						
$\Delta S_{USD,JPY}$	0.64	0.85	0.58					
$\Delta S_{NZD,JPY}$	0.80	0.68	0.61	0.57				
$\Delta S_{NOK,JPY}$	0.63	0.68	0.87	0.59	0.61			
$\Delta S_{SEK,JPY}$	0.62	0.67	0.85	0.57	0.61	0.86		
$\Delta S_{CHF,JPY}$	0.46	0.46	0.83	0.42	0.50	0.76	0.72	
$\Delta S_{GBP,JPY}$	0.58	0.67	0.76	0.63	0.59	0.75	0.73	0.63
Panel III	$\Delta S_{AUD,i}$	$\Delta S_{CAD,i}$	$\Delta S_{EUR,i}$	$\Delta S_{j,i}$	$\Delta S_{NZD,i}$	$\Delta S_{NOK,i}$	$\Delta S_{SEK,i}$	$\Delta S_{CHF,i}$
$\Delta S_{CAD,i}$	-0.22							
$\Delta S_{EUR,i}$	-0.20	-0.32						
$\Delta S_{j,i}$	-0.46	-0.82	-0.13					
$\Delta S_{NZD,i}$	-0.10	-0.24	-0.14	-0.32				
$\Delta S_{NOK,i}$	-0.19	-0.27	-0.04	-0.21	-0.15			
$\Delta S_{SEK,i}$	-0.19	-0.27	-0.04	-0.20	-0.14	-0.04		
$\Delta S_{CHF,i}$	-0.14	-0.27	0.01	0.11	-0.08	-0.01	-0.02	
$\Delta S_{GBP,i}$	-0.24	-0.36	-0.08	-0.27	-0.17	-0.09	-0.09	-0.03

The table presents correlation matrices for all currency pairs relative to the U.S. dollar and Japanese yen in Panels I and II respectively. Panel III reports correlations in Panel I minus those in Panel II. The sample extends from 01/1973 to 12/2015.

Table 2 – Descriptive Statistics of Currency Baskets.

	CB _{USD}	CB _{AUD}	CB _{CAD}	CB _{CHF}	CB _{EUR}	CB _{JPY}	CB _{NOK}	CB _{SEK}	CB _{NZD}	CB _{GBP}
Mean	0.24%	-1.16%	-0.62%	3.67%	-0.83%	2.61%	-0.51%	-1.25%	-1.20%	-0.96%
S.D.	7.92%	9.63%	7.63%	8.34%	5.83%	10.48%	6.42%	6.89%	9.78%	7.40%
AC(1)	0.053	0.017	-0.023	-0.005	0.014	0.050	-0.054	0.100	-0.044	0.033

The table presents annualized means, standard deviations, and the first-order autocorrelation coefficient of currency baskets (CB). The mean is annualized by multiplying by 12. The standard deviation is annualized by multiplying by $\sqrt{12}$. The CB factors are equally-weighted average log appreciations of the indicated G10 currencies relative to the other currencies. The sample extends from 01/1973 to 12/2015.

Table 3 - Explanatory Power of Currency-basket Factors.

	Adjusted R-Square										Top 3	Top 5
	USD	AUD	CAD	CHF	EUR	JPY	NOK	SEK	NZD	GBP		
USD rates	55.4 %	11.5 %	23.6 %	16.0 %	16.1 %	10.7 %	14.2 %	13.7 %	10.1 %	6.2 %	71.8 %	81.5 %
AUD rates	8.0 %	65.5 %	11.7 %	17.8 %	17.8 %	9.7 %	10.9 %	10.2 %	20.4 %	9.2 %	79.0 %	87.4 %
CAD rates	23.4 %	16.4 %	53.4 %	21.9 %	18.0 %	10.1 %	11.4 %	11.1 %	10.9 %	7.1 %	69.6 %	81.5 %
CHF rates	13.0 %	15.9 %	21.1 %	58.9 %	17.3 %	10.2 %	9.2 %	8.1 %	10.6 %	5.7 %	70.0 %	79.3 %
EUR rates	13.6 %	15.2 %	16.3 %	21.8 %	38.7 %	10.3 %	13.8 %	13.4 %	11.9 %	8.2 %	56.5 %	74.3 %
JPY rates	7.7 %	11.3 %	10.6 %	8.6 %	8.5 %	69.9 %	10.6 %	9.7 %	9.1 %	4.9 %	83.2 %	88.2 %
NOK rates	13.5 %	13.2 %	13.1 %	15.0 %	14.9 %	10.4 %	43.5 %	14.6 %	11.7 %	7.4 %	63.4 %	78.0 %
SEK rates	13.4 %	12.7 %	12.6 %	13.3 %	13.8 %	10.0 %	14.1 %	47.6 %	10.9 %	6.9 %	65.7 %	79.1 %
NZD rates	8.6 %	20.8 %	8.1 %	12.7 %	13.9 %	8.7 %	11.9 %	9.8 %	66.3 %	7.0 %	80.2 %	87.2 %
GBP rates	10.1 %	12.8 %	11.4 %	11.8 %	10.7 %	8.8 %	9.1 %	8.9 %	10.2 %	51.7 %	69.7 %	79.9 %
All rates	16.7 %	19.5 %	18.2 %	19.8 %	17.0 %	15.9 %	14.9 %	14.7 %	17.2 %	11.4 %	70.9 %	81.7 %
Off-diagonal	12.4 %	14.4 %	14.3 %	15.4 %	14.6 %	9.9 %	11.7 %	11.1 %	11.8 %	6.9 %		

The table presents average adjusted R-squares of regressing bilateral exchange rates on currency-basket factors from different currency perspectives. For example, the fifth row of the second column (23.4%) is the average adjusted R-square of the following regression, $\Delta s_{j,CAD} = a_j + b_j CB_{USD} = e_{j,CAD}$, $j \in \{G10 \text{ currencies}\}$. We report the explanatory power of the best three and five currency basket factors of each, in the columns indicated by Top 3 and Top 5, respectively. The all rates (off diagonal) rows report the average of the columns (excluding the diagonal entry). The sample extends from 01/1973 to 12/2015.

Table 4 - The Minimum-distance Currency Clusters.

	Panel I				Panel II				
Number of Clusters	Optimal Absolute Clusters	Total Distance	Average Correlation		Optimal Sequential Clusters	Total Distance	Average Correlation		Distance Difference
			Within	Across			Within	Across	
10	(CAD) (USD) (AUD) (CHF) (EUR) (JPY) (NOK) (SEK) (NZD) (GBP)	0	- 1.000	-0.104	(CAD) (USD) (AUD) (CHF) (EUR) (JPY) (NOK) (SEK) (NZD) (GBP)	0	- 1.000	-0.104	0
9	(CAD,USD) (AUD) (CHF) (EUR) (JPY) (NOK) (SEK) (NZD) (GBP)	0.961	0.539 0.958	-0.119	(CAD,USD) (AUD) (CHF) (EUR) (JPY) (NOK) (SEK) (NZD) (GBP)	0.961	0.539 0.958	-0.119	0
8	(CHF,EUR) (CAD,USD) (AUD) (JPY) (NOK) (SEK) (NZD) (GBP)	1.954	0.523 0.920	-0.133	(CHF,EUR) (CAD,USD) (AUD) (JPY) (NOK) (SEK) (NZD) (GBP)	1.954	0.523 0.920	-0.133	0
7	(NZD,AUD) (CHF,EUR) (CAD,USD) (JPY) (NOK) (SEK) (GBP)	3.015	0.494 0.883	-0.147	(NZD,AUD) (CHF,EUR) (CAD,USD) (JPY) (NOK) (SEK) (GBP)	3.015	0.494 0.883	-0.147	0
6	(NOK,SEK) (NZD,AUD) (CHF,EUR) (CAD,USD) (JPY) (GBP)	4.096	0.475 0.850	-0.161	(NOK,SEK) (NZD,AUD) (CHF,EUR) (CAD,USD) (JPY) (GBP)	4.096	0.475 0.850	-0.161	0
5	(JPY,GBP) (NOK,SEK) (NZD,AUD) (CHF,EUR) (CAD,USD)	5.582	0.359 0.786	-0.162	(JPY,GBP) (NOK,SEK) (NZD,AUD) (CHF,EUR) (CAD,USD)	5.582	0.359 0.786	-0.162	0
4	(CHF,JPY,GBP) (EUR,NOK,SEK) (NZD,AUD) (CAD,USD)	9.608	0.263 0.672	-0.184	(CHF,EUR,NOK,SEK) (JPY,GBP) (NZD,AUD) (CAD,USD)	10.361	0.325 0.680	-0.212	0.753
3	(CAD,NZD,AUD) (CHF,EUR,NOK,SEK) (USD,JPY,GBP)	14.830	0.224 0.577	-0.224	(CAD,USD,NZD,AUD) (CHF,EUR,NOK,SEK) (JPY,GBP)	15.957	0.231 0.565	-0.240	1.127
2	(CHF,EUR,NOK,SEK,GBP) (CAD,USD,NZD,AUD,JPY)	26.170	0.125 0.416	-0.287	(CHF,EUR,NOK,SEK,JPY,GBP) (CAD,USD,NZD,AUD)	27.546	0.122 0.405	-0.302	1.376
1	(CHF,EUR,NOK,SEK,JPY,GBP ,CAD,USD,NZD,AUD)	66.224	-0.104 0.096	-	(CHF,EUR,NOK,SEK,JPY,GBP, CAD,USD,NZD,AUD)	66.224	-0.104 0.096	-	0

The table presents the optimal clusters of G10 currency baskets. The distances are computed as in Equation (14), which is based on correlations among currency baskets within a cluster. Panel I presents the optimal clusters generated by the absolute clustering technique, where we consider all possible allocations of G10 currency baskets in N clusters ($1 \leq N \leq 10$) and calculate their in-cluster distance as the sum of distances among all members of the cluster. Panel II presents the optimal clusters resulting from Anh *et al.* (2009)'s sequential clustering technique that at each step merges two clusters, while minimizing the total increase in distance. The average correlation “within” simply is the equally weighted average of all “within” the cluster correlations, where the top number excludes the correlation with the own currency and the bottom number includes it (that is, assigns a 1 to the correlation with oneself). The “across” averages all correlations between currencies not in the same cluster.

Table 5 - Correlations between the G10 Currency Basket Factors.

	CB _{USD}	CB _{AUD}	CB _{CAD}	CB _{CHF}	CB _{EUR}	CB _{JPY}	CB _{NOK}	CB _{SEK}	CB _{NZD}	CB _{GBP}
CB _{USD}	1.00									
CB _{AUD}	-0.03	1.00								
CB _{CAD}	0.54	0.27	1.00							
CB _{CHF}	-0.39	-0.46	-0.56	1.00						
CB _{EUR}	-0.39	-0.45	-0.45	0.51	1.00					
CB _{JPY}	0.09	-0.30	-0.25	0.11	-0.13	1.00				
CB _{NOK}	-0.38	-0.30	-0.26	0.24	0.39	-0.26	1.00			
CB _{SEK}	-0.38	-0.28	-0.27	0.16	0.35	-0.25	0.42	1.00		
CB _{NZD}	-0.17	0.44	0.01	-0.31	-0.35	-0.25	-0.32	-0.26	1.00	
CB _{GBP}	-0.07	-0.30	-0.15	-0.03	0.08	-0.10	0.04	0.00	-0.22	1.00
CF	0.53	0.60	0.57	-0.65	-0.74	0.19	-0.66	-0.62	0.48	-0.38
45FPC	-0.30	-0.81	-0.59	0.72	0.64	0.35	0.46	0.41	-0.66	0.26

The top panel presents monthly correlations between the currency baskets. The bottom panel presents monthly correlations between our simple currency factor (CF) and the FPC of the 45 non-overlapping bilateral exchange rates (45FPC) on the one hand, and the currency-baskets on the other hand. The sample extends from 01/1973 to 12/2015.

Table 6 - Explanatory Power of Various Currency Factors.

	Panel I	Panel II		Panel III			Panel IV		
	CF _{abs}	CF ₃₁	CF ₃₂	CF _{abs}	CF _{com}	CF _{TW}	CF ₃₁	CF ₃₂	CF _{TW}
CB _{USD}	0.20 (12.56)	0.38 (17.96)	0.15 (9.15)	-0.03 (-3.09)	0.19 (12.98)	2.34 (43.77)	-0.05 (-3.41)	0.12 (18.84)	2.25 (44.99)
CB _{AUD}	0.27 (12.17)	-0.06 (-3.34)	0.40 (19.06)	0.17 (6.93)	0.27 (10.66)	-0.59 (-4.12)	0.00 (0.04)	0.40 (18.63)	-0.30 (-2.32)
CB _{CAD}	0.21 (15.14)	0.14 (5.88)	0.25 (13.53)	-0.05 (-2.01)	0.35 (11.59)	1.37 (10.81)	-0.09 (-2.92)	0.24 (15.31)	1.22 (10.39)
CB _{CHF}	-0.26 (-18.55)	-0.24 (-8.86)	-0.33 (-19.84)	-0.08 (-2.99)	-0.27 (-7.52)	-0.72 (-4.75)	-0.22 (-5.60)	-0.33 (-19.83)	-0.10 (-0.74)
CB _{EUR}	-0.21 (-24.48)	-0.22 (-13.10)	-0.24 (-23.14)	-0.24 (-14.97)	0.04 (1.78)	0.34 (3.75)	-0.30 (-14.72)	-0.24 (-24.45)	0.45 (6.18)
CB _{JPY}	0.10 (3.80)	0.47 (16.15)	-0.07 (-3.76)	0.66 (29.27)	-0.92 (-34.02)	-1.79 (-16.27)	0.67 (13.65)	-0.06 (-3.03)	-1.08 (-6.41)
CB _{NOK}	-0.20 (-17.83)	-0.26 (-11.27)	-0.21 (-17.14)	-0.37 (-19.55)	0.28 (12.56)	0.55 (6.31)	-0.24 (-7.26)	-0.21 (-17.06)	-0.13 (-1.40)
CB _{SEK}	-0.20 (-13.13)	-0.28 (-11.08)	-0.22 (-12.75)	-0.18 (-7.22)	0.00 (0.05)	-0.34 (-1.85)	-0.24 (-7.89)	-0.21 (-12.97)	-0.21 (-1.58)
CB _{NZD}	0.22 (10.94)	-0.08 (-3.48)	0.35 (14.44)	0.25 (8.36)	0.10 (3.74)	-1.32 (-7.01)	0.09 (2.78)	0.36 (15.01)	-0.93 (-5.65)
CB _{GBP}	-0.13 (-7.61)	0.16 (5.09)	-0.08 (-4.08)	-0.12 (-3.91)	-0.04 (-1.26)	0.18 (1.23)	0.38 (9.34)	-0.07 (-3.87)	-1.16 (-8.08)
Adj. R ²	0.32 [0.14,0.43]	0.48 [0.35,0.59]		0.58 [0.41,0.81]			0.58 [0.46,0.70]		

The table presents results of regressing G10 currency baskets on our suggested currency factors. Our suggested currency factors include a simple factor based on two absolute clusters as in Equation (15), CF_{abs}, a commodity currency factor as in Equation (20), CF_{com}, a currency trading-volume weighted factor as in Equation (21), CF_{TW}, and two currency factors based on three absolute clusters, CF₃₁ and CF₃₂ as in Equations (18) and (19), respectively. We combine these factors in one univariate (Panel I) and three multivariate models (Panels II through IV). The t-statistics use White standard errors and are reported in parentheses. The table also reports the average R-squares as well as their ranges in brackets. The sample extends from 01/1973 to 12/2015.

Table 7 - Explanatory Power of Extant Currency Factors for Currency Basket Factors.

	Panel I					
	Carry	Commodity	Value	Volatility	Momentum	Equity
CB _{USD}	-0.06 (-0.79)	-0.25 (-6.21)	0.31 (4.13)	0.18 (1.51)	0.05 (0.60)	-0.08 (-2.51)
CB _{AUD}	0.69 (10.46)	0.20 (4.33)	0.50 (6.78)	-0.31 (-2.45)	-0.08 (-0.78)	0.21 (6.49)
CB _{CAD}	0.35 (6.13)	0.02 (0.40)	0.39 (5.70)	-0.12 (-1.27)	-0.09 (-1.38)	0.11 (4.41)
CB _{CHF}	-0.61 (-9.89)	-0.04 (-1.12)	-0.66 (-9.64)	0.17 (1.52)	0.23 (3.16)	-0.16 (-6.62)
CB _{EUR}	-0.17 (-3.19)	0.01 (0.62)	-0.21 (-4.16)	0.06 (0.83)	0.02 (0.31)	-0.09 (-4.46)
CB _{JPY}	-0.86 (-12.41)	-0.25 (-3.06)	-0.28 (-3.23)	0.63 (3.73)	0.12 (0.99)	-0.16 (-3.60)
CB _{NOK}	-0.09 (-1.49)	0.11 (3.92)	-0.53 (-11.41)	-0.17 (-2.05)	-0.05 (-0.76)	0.01 (0.28)
CB _{SEK}	0.00 (0.00)	0.10 (3.86)	-0.35 (-5.97)	-0.18 (-2.66)	-0.05 (-0.89)	0.03 (0.97)
CB _{NZD}	0.63 (8.61)	0.10 (2.08)	0.71 (7.90)	-0.14 (-1.08)	-0.03 (-0.30)	0.17 (4.68)
CB _{GBP}	0.11 (1.73)	-0.01 (-0.42)	0.11 (1.64)	-0.13 (-1.19)	-0.12 (-1.88)	-0.04 (-1.62)
Adj. R ²	0.13 [0.01,0.28]	0.03 [0.00,0.06]	0.12 [0.03,0.24]	0.02 [0.00,0.02]	0.01 [0.00,0.01]	0.04 [0.00,0.07]
	Panel II					
	Carry	Commodity	Value	Volatility	Momentum	Equity
CB _{USD}	-0.13 (-1.77)	-0.20 (-4.97)	0.37 (4.67)	-0.07 (-0.72)	-0.01 (-0.13)	-0.04 (-1.69)
CB _{AUD}	0.51 (5.84)	0.10 (2.62)	0.21 (2.75)	0.01 (0.12)	0.05 (0.71)	0.09 (2.85)
CB _{CAD}	0.19 (2.44)	-0.03 (-0.95)	0.27 (3.26)	-0.06 (-0.58)	-0.05 (-0.73)	0.06 (2.70)
CB _{CHF}	-0.36 (-5.47)	0.04 (1.23)	-0.43 (-6.56)	0.06 (0.54)	0.16 (2.31)	-0.07 (-2.82)
CB _{EUR}	-0.06 (-0.97)	0.05 (1.78)	-0.16 (-2.90)	0.04 (0.49)	0.00 (0.07)	-0.08 (-3.78)
CB _{JPY}	-0.89 (-11.79)	-0.08 (-1.96)	0.17 (2.22)	0.16 (1.57)	-0.04 (-0.64)	0.02 (0.56)
CB _{NOK}	0.14 (2.80)	0.05 (2.30)	-0.60 (-11.79)	-0.01 (-0.08)	-0.04 (-0.88)	0.00 (0.09)
CB _{SEK}	0.14 (1.75)	0.05 (1.85)	-0.42 (-5.83)	-0.03 (-0.38)	-0.04 (-0.67)	0.01 (0.39)
CB _{NZD}	0.37 (3.18)	0.05 (1.14)	0.50 (3.83)	0.04 (0.43)	0.07 (1.12)	0.07 (1.87)
CB _{GBP}	0.08 (1.21)	-0.03 (-0.87)	0.08 (1.07)	-0.15 (-1.37)	-0.11 (-1.76)	-0.06 (-2.45)
Adj. R ²	0.23 [0.10,0.37]					

The table presents results of regressing currency basket returns on the extant currency factors. Panel I reports results of univariate regressions and Panel II reports the results of multivariate regressions. The t-statistics are based on White standard errors and reported in parentheses. The table also reports the average R-squares as well as their ranges in brackets. The sample extends from 01/1973 to 12/2015.

Table 8 - Correlations among Currency Factors as well as Top Three Principal Components.

	Panel I			Panel II									
	45FPC	45SPC	45TPC	CF _{abs}	CF _{com}	CF _{TW}	CF ₃₁	CF ₃₂	Carry	Volatility	Commodity	Momentum	Value
CF _{abs}	-0.83	0.53	0.16										
CF _{com}	-0.90	-0.31	0.05	0.60									
CF _{TW}	0.03	0.77	-0.41	0.32	-0.36								
CF ₃₁	0.22	0.87	-0.09	0.23	-0.54	0.74							
CF ₃₂	-0.97	-0.12	0.13	0.77	0.94	-0.21	-0.37						
Carry	-0.62	-0.40	-0.18	0.26	0.65	-0.28	-0.39	0.62					
Volatility	0.12	0.25	0.13	0.05	-0.19	0.19	0.21	-0.14	-0.29				
Commodity	-0.11	-0.33	0.02	-0.08	0.25	-0.37	-0.35	0.17	0.23	-0.27			
Momentum	0.08	0.06	0.10	-0.01	-0.10	0.06	0.02	-0.07	-0.16	-0.01	-0.15		
Value	-0.56	0.08	-0.01	0.49	0.38	0.12	0.06	0.52	0.47	0.09	-0.07	-0.05	
Equity	-0.33	-0.19	0.00	0.17	0.40	-0.28	-0.26	0.36	0.36	-0.16	0.24	-0.04	0.11

The table presents monthly correlations between the various currency factors considered in this article and the First, Second, and Third Principal Components (FPC, SPC, and TPC respectively) of the 45 non-overlapping bilateral exchange rates. The sample extends from 01/1973 to 12/2015.

Table 9 - Horse Race Part I: Univariate Factors.

	RMSE	BS-RMSE	SE	95% Conf. Interval	
CF _{abs}	0.176	0.187	0.012	0.163	0.211
CF _{com}	0.231	0.226	0.013	0.201	0.252
CF _{TW}	0.281	0.265	0.022	0.222	0.309
CF ₃₁	0.299	0.284	0.022	0.241	0.327
CF ₃₂	0.201	0.202	0.011	0.182	0.223
Carry	0.264	0.254	0.020	0.214	0.294
Volatility	0.298	0.284	0.020	0.245	0.324
Commodity	0.235	0.224	0.016	0.192	0.256
Momentum	0.301	0.287	0.020	0.248	0.325
Value	0.303	0.288	0.019	0.251	0.326
Equity	0.288	0.274	0.020	0.236	0.312

This table presents the RMSEs and the bootstrap results for the RMSE of implied correlations for various univariate models with factors tested in Tables 6 and 7. We bootstrap the G10 currency baskets simultaneously with replacement. For each random sample, we estimate the correlation matrix as well as the factor model. Then, we use the factors exposures to compute model-implied correlations and finally the RMSEs. We use a bock-bootstrap using 6 month-blocks creating samples of the same size as the actual sample. The number of replications is 1000. The sample extends from 01/1973 to 12/2015.

Table 10 - Horse Race Part II: Multivariate Models.

	Panel I				
	RMSE	BS-RMSE	SE	95% Conf. Interval	
CF _{abs} + CF _{com} + CF _{TW}	0.112	0.124	0.009	0.106	0.143
CF ₃₁ + CF ₃₂ + CF _{TW}	0.131	0.142	0.009	0.125	0.160
Carry + Volatility+ Commodity + Momentum + Value + Equity	0.214	0.206	0.019	0.169	0.243
Carry + Commodity + Value + Equity	0.215	0.207	0.019	0.171	0.244
Carry + Commodity + Value	0.218	0.210	0.019	0.172	0.248
Carry + Value + Equity	0.218	0.209	0.018	0.173	0.244
	Panel II				
	RMSE	BS-RMSE	SE	95% Conf. Interval	
CF _{abs} + CF _{com} + CF _{TW} + Carry	0.103	0.115	0.009	0.098	0.132
CF _{abs} + CF _{com} + CF _{TW} + Value	0.104	0.117	0.010	0.098	0.136
CF _{abs} + CF _{com} + CF _{TW} + Carry + Value	0.100	0.112	0.009	0.094	0.129
CF ₃₁ + CF ₃₂ + CF _{TW} + Carry	0.110	0.119	0.009	0.102	0.136
CF ₃₁ + CF ₃₂ + CF _{TW} + Value	0.126	0.137	0.010	0.118	0.155
CF ₃₁ + CF ₃₂ + CF _{TW} + Carry + Value	0.105	0.114	0.009	0.097	0.130

This table presents the RMSEs and the bootstrap results for the RMSE of implied correlations for various currency multivariate models with factors tested in previous tables. Panel I reports models included either extant or new currency factors. Panel II reports some models included both extant and new currency factors together. We bootstrap the G10 currency baskets simultaneously with replacement. For each random sample, we estimate the correlation matrix as well as the factor model. Then, we use the factors exposures to compute model-implied correlations and finally the RMSEs. We use a bock-bootstrap using 6 month-blocks creating samples of the same size as the actual sample. The number of replications is 1000. The sample extends from 01/1973 to 12/2015.

Table 11 - Revisiting Verdelhan (2016) - Explanatory Regression of Bilateral Exchange Rates versus Currency Basket Returns.

	Panel I				Panel II		
	Carry	CB _{USD}	R ²		Carry	CB _{USD}	R ²
$\Delta s_{AUD,USD}$	-0.62 (-13.48)	0.91 (21.82)	0.57	CB _{AUD}	0.69 (13.48)	-0.01 (-0.15)	0.26
$\Delta s_{CAD,USD}$	-0.34 (-10.57)	0.42 (14.25)	0.39	CB _{CAD}	0.38 (10.57)	0.54 (16.50)	0.41
$\Delta s_{CHF,USD}$	0.58 (16.60)	1.29 (41.47)	0.79	CB _{CHF}	-0.64 (-16.60)	-0.44 (-12.60)	0.44
$\Delta s_{EUR,USD}$	0.17 (5.75)	1.17 (44.76)	0.80	CB _{EUR}	-0.19 (-5.75)	-0.30 (-10.27)	0.20
$\Delta s_{JPY,USD}$	0.77 (16.17)	0.82 (19.33)	0.54	CB _{JPY}	-0.85 (-16.17)	0.08 (1.77)	0.34
$\Delta s_{NOK,USD}$	0.09 (2.85)	1.18 (39.75)	0.75	CB _{NOK}	-0.10 (-2.85)	-0.31 (-9.42)	0.15
$\Delta s_{SEK,USD}$	0.02 (0.48)	1.20 (37.34)	0.73	CB _{SEK}	-0.02 (-0.48)	-0.33 (-9.25)	0.14
$\Delta s_{NZD,USD}$	-0.56 (-11.68)	1.06 (24.66)	0.60	CB _{NZD}	0.62 (11.68)	-0.18 (-3.76)	0.23
$\Delta s_{GBP,USD}$	-0.09 (-2.30)	0.95 (25.84)	0.57	CB _{GBP}	0.10 (2.30)	-0.06 (-1.46)	0.01

The table reports coefficients and R²'s from regressing bilateral exchange rates against the U.S. dollar on the carry and the U.S. basket factor (Panel I) and from regressing currency basket factors on the carry and the U.S. basket factor (Panel II). The t-statistics are based on White standard errors and reported in parentheses. The sample extends from 01/1973 to 12/2015.

Table 12 - Explanatory Power of the Global Dollar Factor (HML\$) of Verdelhan (2016) versus Our Currency Factor (CF).

	Panel I		Panel II		Panel III		
	HML\$	R ²	CF _{abs}	R ²	HML\$	CF _{abs}	R ²
CB _{USD}	0.006 (21.47)	0.71	0.16 (7.11)	0.19	0.007 (17.45)	-0.03 (-1.87)	0.71
CB _{AUD}	0.001 (1.62)	0.03	0.34 (16.80)	0.57	-0.004 (-9.17)	0.45 (20.79)	0.69
CB _{CAD}	0.004 (6.00)	0.27	0.26 (16.92)	0.50	0.001 (1.65)	0.22 (7.44)	0.51
CB _{CHF}	-0.003 (-5.51)	0.23	-0.24 (-15.42)	0.48	-0.001 (-1.09)	-0.22 (-7.81)	0.49
CB _{EUR}	-0.004 (-12.56)	0.47	-0.18 (-16.11)	0.48	-0.002 (-6.35)	-0.12 (-7.96)	0.60
CB _{JPY}	0.002 (2.05)	0.04	-0.14 (-3.77)	0.07	0.006 (6.29)	-0.31 (-7.21)	0.27
CB _{NOK}	-0.003 (-5.76)	0.23	-0.16 (-9.03)	0.28	-0.002 (-2.05)	-0.11 (-3.34)	0.32
CB _{SEK}	-0.004 (-8.82)	0.28	-0.16 (-7.16)	0.23	-0.003 (-6.36)	-0.08 (-3.48)	0.32
CB _{NZD}	0.000 (0.58)	0.00	0.24 (11.58)	0.37	-0.004 (-5.55)	0.36 (10.72)	0.52
CB _{GBP}	0.000 (-0.77)	0.00	-0.11 (-5.63)	0.11	0.001 (1.96)	-0.16 (-4.76)	0.14
Adj. R ² (All)	0.23		0.33		0.46		
Adj. R ² (Non-Dollar)	0.21		0.27		0.35		
RMSE	0.304		0.179		0.266		

The table compares the explanatory power of the global dollar factor (HML\$) of Verdelhan (2016) and our simple currency factor. Panel I reports the results of regressing currency baskets on HML\$. Panel II reports the results of regressing currency basket factors on our currency factor. Panel III reports the results of regressing currency basket factors on HML\$ and our currency factor. The t-statistics are based on White standard errors and reported in parentheses. The RMSE is the Root Mean Squared Error of the implied correlations. The global dollar factor (HML\$) of Verdelhan (2016) is available from 11/1988 to 12/2010.

Table 13 - Revisiting Lustig and Richmond (2016) – The R2 of Bilateral Exchange Rates versus Currency Basket Returns.

	Panel I			Panel II	
	CB _{USD}	R ²		CB _{USD}	R ²
$\Delta s_{AUD,USD}$	0.94 (14.48)	0.42	CB _{AUD}	-0.04 (-0.55)	0.00
$\Delta s_{CAD,USD}$	0.43 (8.48)	0.26	CB _{CAD}	0.52 (9.14)	0.29
$\Delta s_{CHF,USD}$	1.27 (24.16)	0.68	CB _{CHF}	-0.41 (-6.98)	0.15
$\Delta s_{EUR,USD}$	1.16 (39.32)	0.78	CB _{EUR}	-0.29 (-8.82)	0.15
$\Delta s_{JPY,USD}$	0.79 (9.51)	0.31	CB _{JPY}	0.12 (1.34)	0.01
$\Delta s_{NOK,USD}$	1.17 (31.63)	0.75	CB _{NOK}	-0.31 (-7.40)	0.14
$\Delta s_{SEK,USD}$	1.20 (28.12)	0.73	CB _{SEK}	-0.33 (-6.95)	0.14
$\Delta s_{NZD,USD}$	1.09 (16.06)	0.50	CB _{NZD}	-0.21 (-2.77)	0.03
$\Delta s_{GBP,USD}$	0.96 (20.49)	0.57	CB _{GBP}	-0.06 (-1.24)	0.00

The table reports coefficients and R²'s from regressing bilateral exchange rates against the U.S. dollar on the U.S. basket factor (Panel I) and from regressing currency basket factors on the U.S. basket factor (Panel II). The t-statistics are based on White standard errors and reported in parentheses. The sample extends from 01/1973 to 12/2015.

Table 14 – Explaining the Variation in 15 Emerging Currency Basket Changes.

	Panel I	Panel II		Panel III			Panel IV		
	CF _{abs}	CF ₃₁	CF ₃₂	CF _{abs}	CF _{com}	CF _{TW}	CF ₃₁	CF ₃₂	CF _{TW}
CB _{CNY}	0.09 (3.65)	0.29 (12.83)	0.02 (0.76)	-0.10 (-1.61)	0.13 (1.62)	2.03 (4.95)	-0.05 (-0.72)	0.00 (0.07)	1.74 (4.73)
CB _{CZK}	-0.17 (-4.61)	-0.29 (-7.72)	-0.16 (-3.16)	-0.16 (-4.11)	0.04 (0.50)	-0.35 (-1.46)	-0.32 (-3.41)	-0.17 (-3.32)	0.14 (0.38)
CB _{HUF}	-0.16 (-6.89)	-0.30 (-8.09)	-0.13 (-4.75)	-0.20 (-5.13)	0.13 (2.54)	-0.19 (-0.75)	-0.32 (-4.73)	-0.13 (-4.82)	0.10 (0.39)
CB _{INR}	0.10 (4.03)	0.16 (5.55)	0.08 (3.15)	0.01 (0.27)	0.09 (2.14)	0.86 (4.89)	0.00 (-0.07)	0.08 (3.03)	0.82 (4.57)
CB _{IDR}	0.14 (2.01)	0.11 (1.24)	0.13 (1.74)	0.19 (1.03)	-0.07 (-0.38)	-0.33 (-0.40)	0.12 (0.53)	0.13 (1.64)	-0.05 (-0.07)
CB _{MXN}	0.17 (5.87)	0.15 (3.57)	0.22 (6.31)	-0.03 (-0.56)	0.29 (4.63)	1.16 (4.16)	0.00 (0.04)	0.22 (6.10)	0.77 (2.66)
CB _{PHP}	0.15 (6.40)	0.22 (7.19)	0.13 (4.58)	0.07 (1.64)	0.06 (1.34)	0.89 (3.90)	0.03 (0.61)	0.13 (4.66)	0.94 (4.22)
CB _{PLN}	-0.12 (-2.97)	-0.26 (-7.17)	-0.08 (-1.82)	-0.12 (-2.81)	0.09 (1.93)	-0.48 (-2.08)	-0.20 (-2.95)	-0.08 (-1.81)	-0.30 (-0.98)
CB _{SGD}	0.06 (5.15)	0.14 (8.36)	0.02 (1.79)	0.07 (3.71)	-0.07 (-3.02)	0.23 (2.16)	0.10 (3.80)	0.02 (1.67)	0.18 (1.61)
CB _{ZAR}	0.04 (1.25)	-0.08 (-1.87)	0.10 (2.92)	0.01 (0.17)	0.12 (1.72)	-0.33 (-1.07)	0.00 (-0.03)	0.11 (3.05)	-0.41 (-1.41)
CB _{KRW}	0.11 (4.44)	0.06 (1.13)	0.13 (4.07)	0.15 (2.18)	-0.04 (-0.44)	-0.43 (-1.01)	0.10 (0.94)	0.13 (3.87)	-0.22 (-0.51)
CB _{TWD}	0.09 (6.18)	0.20 (10.78)	0.04 (2.39)	0.06 (2.57)	-0.02 (-0.94)	0.71 (5.98)	0.06 (1.95)	0.04 (2.32)	0.71 (5.81)
CB _{THB}	0.09 (4.85)	0.17 (6.48)	0.06 (2.50)	0.10 (2.75)	-0.07 (-1.61)	0.30 (1.41)	0.11 (2.41)	0.06 (2.39)	0.34 (1.79)
CB _{TRY}	0.06 (1.52)	-0.05 (-1.18)	0.13 (2.57)	-0.11 (-1.66)	0.30 (3.81)	0.72 (1.79)	-0.20 (-2.46)	0.12 (2.36)	0.76 (1.96)
Ave. Adj. R ²	0.06 [0.01,0.13]	0.13 [0.03,0.27]		0.15 [0.04,0.32]			0.16 [0.04,0.31]		
RMSE	0.159	0.121		0.121			0.121		
Block-Bootstrap 95% C.I.	[0.154,0.222]	[0.108,0.168]		[0.115,0.179]			[0.108,0.170]		
	Panel V					Panel VI			
	CF _{abs}	CF _{com}	CF _{EM}	CF _{abs}	CF _{com}	CF _{TW}	CF _{EM}		
CB _{CNY}	0.18 (6.83)	-0.25 (-11.05)	0.02 (0.84)	-0.10 (-4.03)	0.16 (4.32)	2.18 (11.60)	-0.06 (-5.41)		
CB _{CZK}	-0.18 (-6.76)	0.11 (1.88)	-0.09 (-3.25)	-0.16 (-10.41)	0.08 (2.43)	-0.15 (-1.54)	-0.08 (-7.41)		
CB _{HUF}	-0.22 (-7.96)	0.16 (5.13)	-0.02 (-0.90)	-0.20 (-12.62)	0.13 (6.25)	-0.15 (-1.23)	-0.02 (-1.44)		
CB _{INR}	0.11 (3.55)	-0.07 (-3.03)	0.06 (2.71)	0.01 (0.71)	0.07 (4.17)	0.77 (10.13)	0.04 (3.49)		
CB _{IDR}	0.14 (1.59)	-0.01 (-0.11)	0.01 (0.12)	0.19 (2.54)	-0.08 (-1.05)	-0.38 (-1.08)	0.02 (0.85)		
CB _{MXN}	0.04 (1.63)	0.07 (2.25)	0.30 (5.42)	-0.02 (-1.17)	0.16 (6.06)	0.49 (3.95)	0.28 (11.48)		
CB _{PHP}	0.19 (7.18)	-0.10 (-4.42)	0.01 (0.52)	0.07 (4.01)	0.07 (3.69)	0.94 (9.76)	-0.02 (-3.02)		
CB _{PLN}	-0.20 (-4.78)	0.18 (5.45)	0.01 (0.45)	-0.12 (-6.95)	0.07 (3.63)	-0.56 (-5.63)	0.03 (2.49)		
CB _{SGD}	0.09 (7.19)	-0.11 (-7.26)	0.02 (1.70)	0.07 (9.11)	-0.08 (-7.97)	0.19 (4.09)	0.02 (2.77)		
CB _{ZAR}	-0.11 (-3.38)	0.18 (4.87)	0.23 (6.19)	0.02 (0.65)	0.00 (-0.02)	-0.97 (-6.83)	0.27 (15.17)		
CB _{KRW}	0.05 (1.44)	0.04 (1.04)	0.13 (4.02)	0.16 (5.76)	-0.11 (-3.59)	-0.82 (-4.75)	0.16 (11.45)		
CB _{TWD}	0.14 (9.35)	-0.16 (-9.60)	0.04 (2.57)	0.06 (6.38)	-0.03 (-2.77)	0.67 (13.42)	0.01 (2.28)		
CB _{THB}	0.14 (6.65)	-0.12 (-5.25)	0.00 (0.07)	0.10 (6.74)	-0.06 (-3.62)	0.32 (3.61)	-0.01 (-1.24)		
CB _{TRY}	-0.02 (-0.56)	0.16 (3.39)	0.05 (0.72)	-0.11 (-4.10)	0.29 (9.26)	0.66 (4.44)	0.02 (0.87)		
Ave. Adj. R ²	0.15 [0.03,0.27]			0.18 [0.04,0.33]					
RMSE	0.128			0.124					
Block-Bootstrap 95% C.I.	[0.122,0.182]			[0.113,0.177]					
	Panel VII					Panel VIII			

	CF ₃₁	CF ₃₂	CF _{EM}	CF ₃₁	CF ₃₂	CF _{TW}	CF _{EM}
CB _{CNY}	0.29 (28.29)	0.02 (1.60)	0.00 (0.46)	-0.05 (-1.88)	0.02 (1.85)	1.85 (11.17)	-0.05 (-5.19)
CB _{CZK}	-0.26 (-15.43)	-0.14 (-7.23)	-0.07 (-6.36)	-0.32 (-8.61)	-0.14 (-7.14)	0.31 (2.20)	-0.08 (-8.12)
CB _{HUF}	-0.29 (-17.72)	-0.12 (-11.73)	-0.01 (-1.30)	-0.32 (-11.70)	-0.12 (-11.76)	0.14 (1.28)	-0.02 (-1.65)
CB _{INR}	0.14 (10.62)	0.06 (5.05)	0.06 (6.37)	0.00 (-0.02)	0.06 (5.33)	0.73 (9.46)	0.04 (3.87)
CB _{IDR}	0.10 (2.62)	0.12 (3.94)	0.02 (0.71)	0.12 (1.32)	0.12 (3.99)	-0.09 (-0.27)	0.02 (0.79)
CB _{MXN}	0.05 (3.46)	0.12 (9.84)	0.29 (13.11)	0.02 (0.73)	0.12 (9.85)	0.17 (1.33)	0.29 (12.13)
CB _{PHP}	0.22 (16.61)	0.13 (10.38)	0.01 (0.72)	0.03 (1.44)	0.13 (11.51)	0.99 (10.76)	-0.02 (-2.98)
CB _{PLN}	-0.27 (-16.92)	-0.09 (-4.63)	0.03 (2.11)	-0.20 (-7.14)	-0.09 (-4.66)	-0.37 (-2.83)	0.04 (2.76)
CB _{SGD}	0.13 (18.09)	0.01 (2.80)	0.02 (3.54)	0.10 (9.52)	0.02 (2.90)	0.15 (3.15)	0.02 (2.64)
CB _{ZAR}	-0.16 (-10.24)	0.02 (1.43)	0.24 (15.46)	0.02 (0.54)	0.02 (1.35)	-0.97 (-7.24)	0.27 (15.44)
CB _{KRW}	0.01 (0.34)	0.08 (5.56)	0.14 (10.38)	0.11 (2.71)	0.08 (5.33)	-0.55 (-3.14)	0.16 (11.15)
CB _{TWD}	0.19 (25.20)	0.03 (4.05)	0.03 (6.02)	0.06 (4.87)	0.03 (4.84)	0.68 (13.15)	0.01 (2.51)
CB _{THB}	0.17 (15.45)	0.06 (6.05)	0.00 (-0.14)	0.11 (5.91)	0.06 (6.28)	0.36 (4.65)	-0.01 (-1.49)
CB _{TRY}	-0.07 (-3.73)	0.11 (4.83)	0.05 (1.92)	-0.20 (-6.02)	0.11 (4.79)	0.71 (4.62)	0.03 (1.09)
Ave. Adj. R ²	0.16 [0.04,0.28]			0.18 [0.04,0.32]			
RMSE	0.125			0.124			
Block-Bootstrap 95% C.I.	[0.107,0.164]			[0.106,0.168]			

The table presents results of regressing 14 emerging currency basket returns on our suggested currency factors. The set of emerging countries includes China, Czech Republic, Hungary, India, Indonesia, Mexico, Philippines, Poland, Singapore, South Africa, South Korea, Taiwan, Thailand, and Turkey. The emerging currency baskets are the average appreciation rate of the emerging currency w.r.t. 23 currencies; 13 other emerging currencies and the G10 currencies. The different panels correspond to different factor models. Our suggested currency factors include a simple factor based on two absolute clusters as in equation (16), CF_{abs}, two currency factors based on three absolute clusters as in equations (19) and (20), CF₃₁ and CF₃₂, a commodity currency factors as in equation (21), CF_{com}, a currency trading-volume weighted factor as in equation (22), CF_{TW}, and an emerging currency trading-volume weighted factor as in equation (22), CF_{EM}, where the weights are from the BIS and reported in the online Appendix. The t-statistics are based on White standard errors and reported in parentheses. The table also reports the average R-squares as well as their ranges in brackets. The RMSE is the room mean square estimation error of implied correlations. The sample extends from 01/1990 to 12/2015.